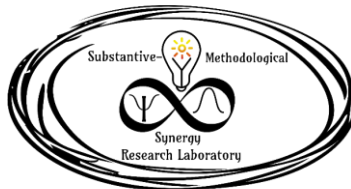


Department of National Defense Longitudinal Analysis Training
Seminar

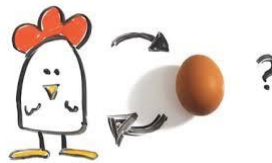


Presented by: Simon Houle

In collaboration with Dr. Alexandre Morin and the Substantive
Methodological Synergy Research Laboratory

Why Longitudinal Data?

- Directionality



- Change



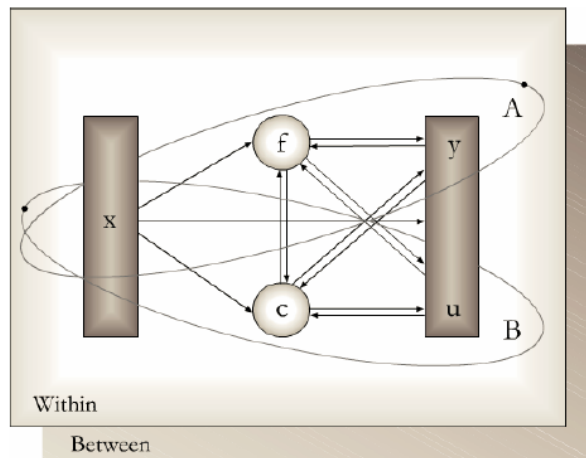
- Shape of change

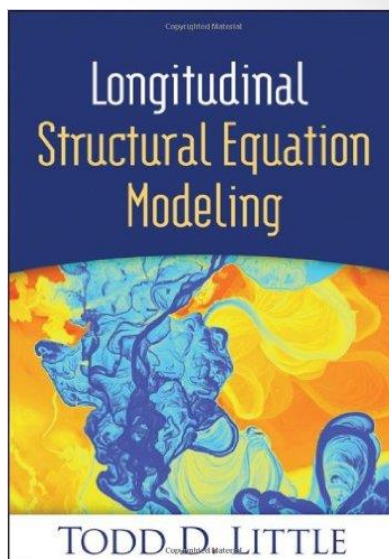
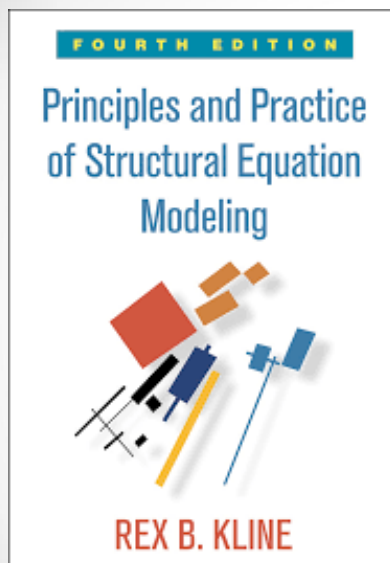
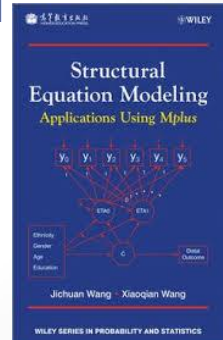
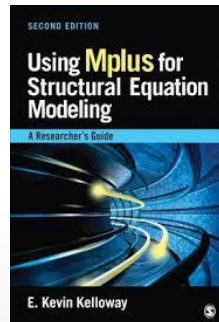
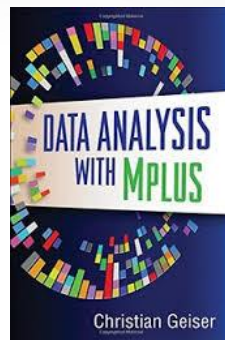
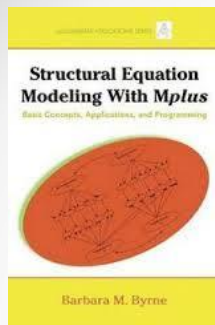


Overview

- Brief introduction to *Mplus*
- Setting the scale for latent variables
- The meaning of time
- Longitudinal measurement invariance in *Mplus*
- Basic longitudinal models
- Auto-regressive crossed-lagged models
- State-trait-error models
- Latent change models
- Cross sectional application for latent change models
- Latent curve models

Mplus





Muthén & Muthén, Mplus X

https://www.statmodel.com

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Mplus
 Tuesday September 13, 2016
 Last updated: August 25, 2016

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 FAQ

MPLUS DEMO VERSION

TRAINING
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 and Handouts
 Web Training

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 Mplus Web Notes
 User's Guide Examples
 Mplus Book
 Mplus Book Examples
 Mplus Book Errata

ANALYSES/RESEARCH
 Mplus Examples
 Papers
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SPECIAL MPLUS TOPICS
 BSEM (Bayesian SEM)
 Complex Survey Data
 SEM (Explanatory SEM)
 Genetics
 IRT
 Measurement Invariance
 Mediation Analysis
 Missing Data
 Mixture Modeling
 Multilevel Modeling
 Structural Equation Modeling
 Survival Analysis
 Randomized Trials

How-To
 Using Mplus via R

Latest News

- Announcing our new book [Regression And Mediation Analysis Using Mplus](#). 140 Mplus inputs, outputs, and data sets where available are posted for the book [examples](#). The book can be ordered from the [online store](#).
- [Mplus favored for Bayesian analysis](#).
- [Mplus Version 7.4](#) is now available. The [Version 7.4 Mplus Language Addendum](#) is found on the website along with the revised [Mplus Version 7 User's Guide](#). Registered users who purchased Mplus within the last year and those with a current Mplus Upgrade and Support Contract can download Version 7.4 using our [online system](#) at no cost.

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Mplus Demo Version

The Mplus Demo version is available for download at no cost. Click [here](#) to download the demo. The demo version contains all of the capabilities of the regular version of Mplus and is only limited by the number of observed variables that can be used in an analysis.

Student Pricing for Mplus Version 7.4

Special student pricing is available for Mplus. The student version of the program is identical to the regular version. Click [here](#) for more information.

Mplus Version 7 User's Guide and Examples

Click [here](#) for the Mplus Version 7 User's Guide and to download the input, output, and data for the Mplus User's Guide examples.

Mplus Web Training and Handouts

Videos and handouts for the 9 topics of the [Mplus Short Courses](#) are now available for viewing on the web. Other [Mplus web training](#) includes web talks, a seminar series, a one-day overview course, a two-day course, and a 20-lecture course on Mplus analyses.

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Mplus User's Guide and Diagrammer Documentation

[Version 7 Mplus User's Guide \(PDF\)](#)
[Version 7 Mplus User's Guide \(HTML\)](#)
[Version 7.1 Examples](#)
[Version 7.4 Mplus Language Addendum](#)
[Version 1 Mplus Diagrammer Documentation](#)

Mplus User's Guide Examples

Following are excerpts from the Version 6 Mplus User's Guide. Chapters 3 - 12 include over 150 examples. These examples are also included on the Mplus CD along with the corresponding Monte Carlo simulation setups that generated the data.

[Chapter 1: Introduction](#)
[Chapter 2: Getting started with Mplus](#)
[Chapter 3: Regression and path analysis](#) [view examples](#)
[Chapter 4: Exploratory factor analysis](#) [view examples](#)
[Chapter 5: Confirmatory factor analysis and structural equation modeling](#) [view examples](#)
[Chapter 6: Growth modeling and survival analysis](#) [view examples](#)
[Chapter 7: Mixture modeling with cross-sectional data](#) [view examples](#)
[Chapter 8: Mixture modeling with longitudinal data](#) [view examples](#)
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TUESDAY SEPTEMBER 13, 2016

Videos and Handouts for Mplus Short Courses

Topic 1. Introductory - advanced factor analysis and structural equation modeling with continuous outcomes.

- Recorded presentation at Johns Hopkins University, August 20, 2009: [768k](#) or [256k](#) bitrate ([handout](#) used in video)
- This video is also available for download in AVI format. Muthen & Muthen holds the copyright for the Mplus Short Course videos. Unauthorized copying of Mplus Short Course Videos is expressly forbidden. The videos are available for personal use and cannot be altered or sold.

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- Most recent [handout](#) of 8/6/09

Topic 2. Introductory - advanced regression analysis, IRT, factor analysis and structural equation modeling with categorical, censored, and count outcomes.

- Recorded presentation at Johns Hopkins University, August 21, 2009: [768k](#) or [256k](#) bitrate ([handout](#) used in video)
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References on this page are ordered by topic. References can also be viewed [ordered by date](#).

Bayesian Analysis [expand topic](#)

Categorical Factor Analysis [expand topic](#)

Complex Survey Data Analysis [expand topic](#)

Dyadic Analysis [expand topic](#)

Exploratory Structural Equation Modeling (ESEM) [expand topic](#)

Factor Mixture Analysis [expand topic](#)

General Mixture Modeling [expand topic](#)

Genetics Modeling [expand topic](#)

Growth Mixture Modeling [expand topic](#)

Growth Modeling [expand topic](#)

IRT [expand topic](#)

Latent Class Analysis [expand topic](#)

Latent Transition Analysis [expand topic](#)

Measurement Modeling [expand topic](#)

Mediation Modeling [expand topic](#)

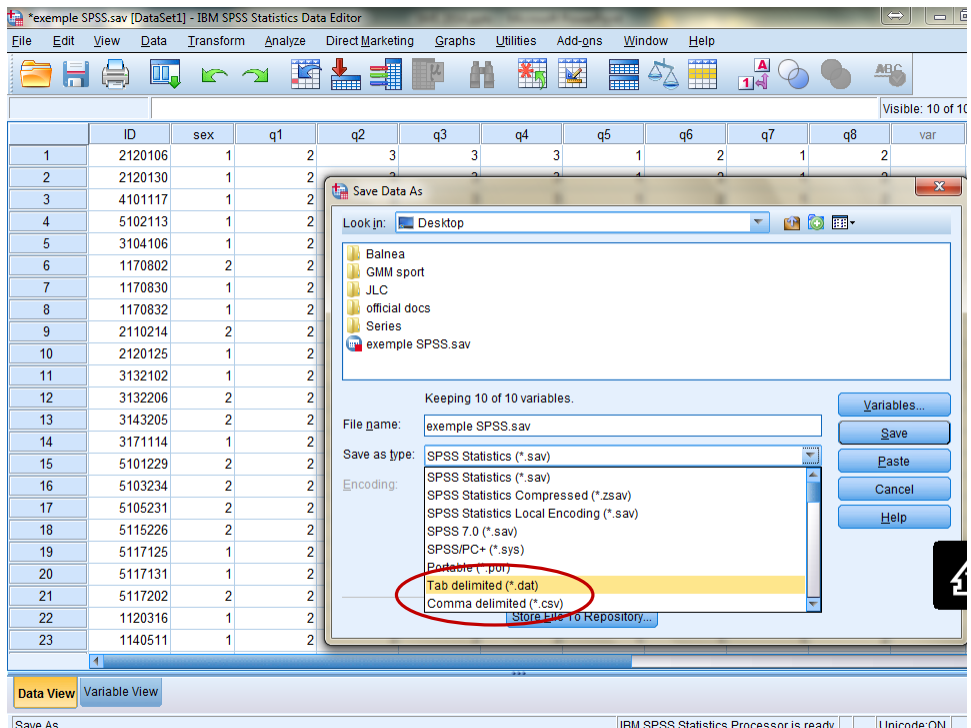
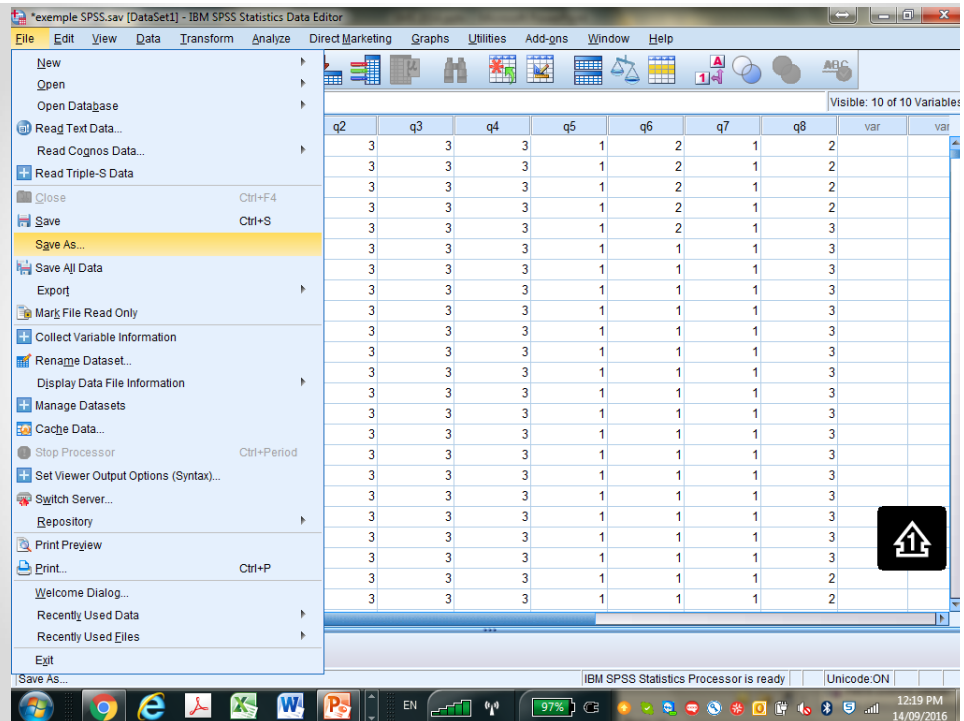
Missing Data Analysis [expand topic](#)

Mixture (Latent Class) SEM [expand topic](#)

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Title:**Title of the model to be estimated;**

! Annotations following “!” are discarded by the program

! The TITLE function is not mandatory.

! All commands end with “;”

! All section titles end with “:”

Data:**File is longitudinaldata.csv;**

! The FILE function of the DATA section is used to identify your

! data set. If the data set is in the same folder, then this is fine.

! If the data set is in another folder, then the full link is indicated.

**File is D:\DOCUMENTS\Cours\LATENT VARIABLE
MODELING\longitudinaldata.csv;**

VARIABLE:**Names are ID sex q1 q2 q3 q4 q5 q6 q7 q8;**

! The NAMES function lists, in order, all variables in the data set.

Usevariables are q1 q2 q3 q4 q5 q6 q7 q8;

! The USEVARIABLES function lists those used in the analysis.

Missing are all (-999);

! The MISSING function identifies the missing data code.

Idvariable = ID;

! The IDVARIABLE identifies the unique identifier.

Auxiliary = sex;**Auxiliary = sex (m);**

! Sometimes, one wants to save the results from an analysis to an

! external data file (e.g., scores on the factors). This external data

! file will include all variables included in the analyses + those

! listed as auxiliary. The (m) indicators allows auxiliary variables

! to be taken into account in the missing data process.

ANALYSIS:

TYPE = General;

ESTIMATOR = MLR; ! Or ML, etc.

! MLR is robust to multivariate non-normality

! MLR can be made to be robust to nesting.

MODEL:

!!! This is where everything happens !



A simple correction for nesting:

VARIABLE:

NAMES = ID sex q1 q2 q3 q4 q5 q6 q7 q8;

USEVARIABLES = q1 q2 q3 q4 q5 q6 q7 q8;

Missing are all (-999);

Idvariable = ID;

CLUSTER = Unit;

Analysis:

TYPE = **COMPLEX;**

ESTIMATOR = MLR;

MODEL:

!!! This is where everything happens !

OUTPUT:

SAMPSTAT **STANDARDIZED** **MODINDICES** **CINTERVAL**
RESIDUAL **SVALUES** **TECH1** **TECH3** **TECH4** ;

SAMPSTAT: sample descriptive.

STANDARDIZED: Standardized parameter estimates.

CINTERVAL: Confidence intervals for parameter estimates.

RESIDUAL: Residuals for parameter estimates.

MODINDICES: Modification indices.

SVALUES: Starts Values.

TECH1: Parameter specifications and starts values (not for EFA).

TECH3: Correlations and covariances for parameter estimates.

TECH4: Means, Correlations and covariances for the latent variables.

MODEL:

ON: Defines a regression e.g., **Y ON X**;

WITH: Defines a correlation e.g., **X WITH Y**;

BY: Defines a factor loading e.g., **FI BY XI X2 X3**;

[]: Variable names within brackets define intercepts and means e.g., **[XI]**; or **[FI]**;

Variable names: By themselves, variable names define variances, uniquenesses and disturbances e.g., **XI**; or **FI**;

*****: Is used to request the free estimation of a parameter that would otherwise be constrained e.g., **FI BY XI* X2 X3**; or to provide a start value for a parameter e.g., **FI BY XI*.900 X2*.850 X3*.800**;

@: Is used to constrain a parameter to a specific value e.g., **FI BY XI@1 X2 X3**;

(): alphanumeric codes in parentheses following a parameter can be used to constrain parameters to equality, e.g. **FI BY XI* (11) X2 (12)**;

***WARNING

Data set contains cases with missing on x-variables.

These cases were not included in the analysis.

Number of cases with missing on x-variables: 61

***WARNING

Data set contains cases with missing on all variables except x-variables. These cases were not included in the analysis.

Number of cases with missing on all variables except x-variables: 30

By explicitly requesting the free
Estimation of the variance of the
Exogenous variables in the MODEL:
section, FIML will be activated.
X1 X2 X3;

THE MODEL ESTIMATION TERMINATED NORMALLY

MODEL FIT INFORMATION

Number of Free Parameters 13

Loglikelihood

H0 Value -7221.664

H0 Scaling Correction Factor 1.6859
for MLR

H1 Value -7221.604

H1 Scaling Correction Factor 1.6256
for MLR

Information Criteria

Akaike (AIC) 14469.328

Bayesian (BIC) 14537.107

Sample-Size Adjusted BIC 14495.811

($n^* = (n + 2) / 24$)

Chi-Square Test of Model Fit

Value	97.470*
Degrees of Freedom	19
P-Value	0.0000
Scaling Correction Factor for MLR	1.4070

* The chi-square value for MLM, MLMV, MLR, ULSMV, WLSM and WLSMV cannot be used for chi-square difference testing in the regular way. MLM, MLR and WLSM chi-square difference testing is described on the Mplus website. MLMV, WLSMV, and ULSMV difference testing is done using the DIFFTEST option.

RMSEA (Root Mean Square Error Of Approximation)

Estimate	0.055
90 Percent C.I.	0.045 0.066
Probability RMSEA <= .05	0.203

CFI/TLI

CFI	0.947
TLI	0.922

SRMR (Standardized Root Mean Square Residual)

Value	0.040
-------	-------

$$\Delta\chi^2$$

The difference in the χ^2 of two nested models is distributed as a χ^2 with degrees of freedom corresponding to the difference in degrees of freedom between the two models.

BUT

* The chi-square value for MLM, MLMV, MLR, ULSMV, WLSM and WLSMV cannot be used for chi-square difference testing in the regular way. MLM, MLR and WLSM chi-square difference testing is described on the Mplus website. MLMV, WLSMV, and ULSMV difference testing is done using the DIFFTEST option.

For MLR, one needs to rely on the Satorra-Bentler correction.

<https://www.statmodel.com/chidiff.shtml>

MODEL RESULTS

					Two-Tailed	
		Estimate	S.E.	Est./S.E.	P-Value	
POSF	BY					Unstandardized loading
Q1		1.000	0.000	999.000	999.000	
Q2		1.175	0.061	19.371	0.000	
POSF	ON					Unstandardized regression (b)
ZSELFEST		0.225	0.027	8.272	0.000	
ZDEPRESS	ON					
ZSELFEST		-0.346	0.037	-9.418	0.000	Covariance
ZSELFEST	WITH					
ZLONELY		-0.246	0.032	-7.688	0.000	
Means						Means, Intercept, Variances
Intercepts						
Variances						
Residual Variances						Residuals (Disturbances, uniquenesses)
POSF		0.936	0.050	18.751	0.000	
Q1		0.883	0.049	17.982		

STANDARDIZED MODEL RESULTS

STDYX Standardization

		Estimate	S.E.	Est./S.E.	P-Value	Two-Tailed
POSF	BY					Standardized loading
Q1		0.583	0.028	21.101	0.000	
Q2		0.727	0.023	31.835	0.000	
POSF	ON					Standardized regression (β)
ZSELFEST		0.225	0.027	8.472	0.000	
ZDEPRESS	ON					
ZSELFEST		-0.346	0.034	-10.305	0.000	
ZSELFEST	WITH					Correlation (r)
ZLONELY		-0.246	0.028	-8.875	0.000	
Means						NA
Intercepts						
Variances						
Residual Variances						Standardized Residuals
POSF		0.936	0.013	72.868	0.000	
Q1		0.882	0.022	39.564	0.000	

R-SQUARE

Observed

Variable

Estimate

S.E.

Est./S.E.

% Explained
varianceCommunality (h^2)

ZGPA

0.064

0.013

4.948

ZDEPRESS

0.118

0.022

5.297

Q1

0.340

0.032

10.550

Q2

0.529

0.033

15.917

0.000

Q3

0.430

0.035

12.212

0.000

Q4

0.551

0.035

15.578

0.000

Q5

0.753

0.071

10.544

0.000

CONFIDENCE INTERVALS OF MODEL RESULTS

Lower .5% Lower 2.5% Lower 5% Estimate Upper 5% Upper 2.5% Upper .5%

ZGPA ON

ZSELFEST

0.155

0.172

0.180

0.225

0.270

0.278

0.295

ZDEPRESS

-0.141

-0.122

-0.112

-0.060

-0.009

0.001

0.021

ZDEPRESS ON

ZSELFEST

-0.441

-0.418

-0.407

-0.346

-0.286

-0.274

-0.252

ZLONELY

-0.087

-0.068

-0.059

-0.009

0.041

0.050

0.069

[...]

CONFIDENCE INTERVALS OF STANDARDIZED MODEL RESULTS**STDYX Standardization**

Lower .5% Lower 2.5% Lower 5% Estimate Upper 5% Upper 2.5% Upper .5%

ZGPA ON

ZSELFEST

0.157

0.173

0.181

0.225

0.269

0.277

0.293

ZDEPRESS

-0.142

-0.122

-0.112

-0.060

-0.008

0.001

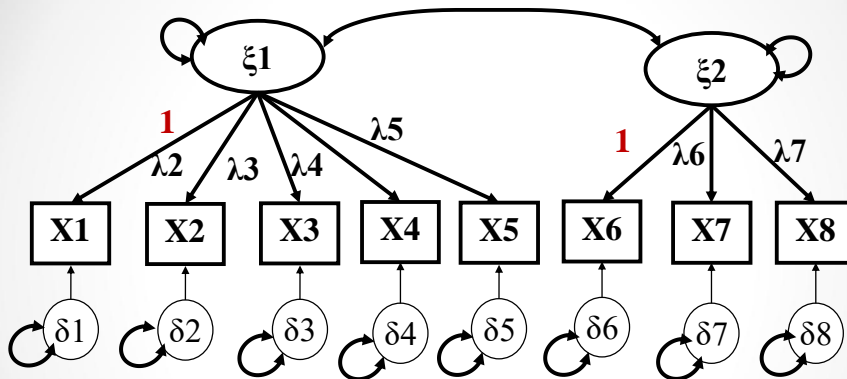
0.021

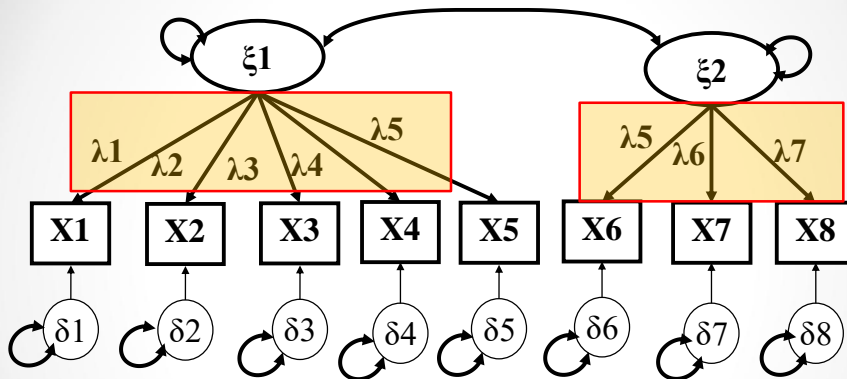
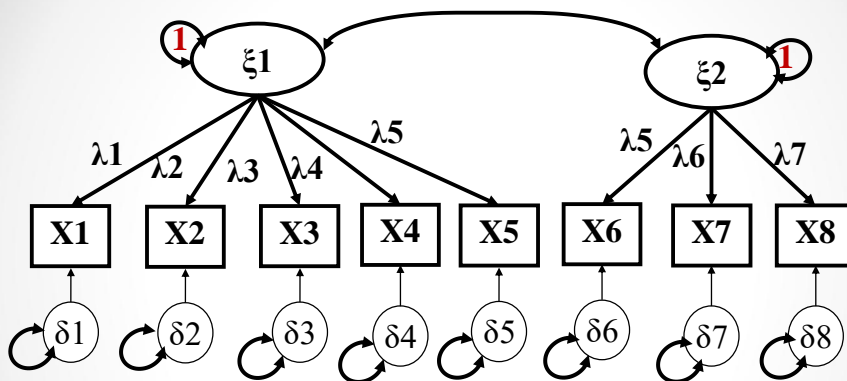
Preliminary Measurement Models

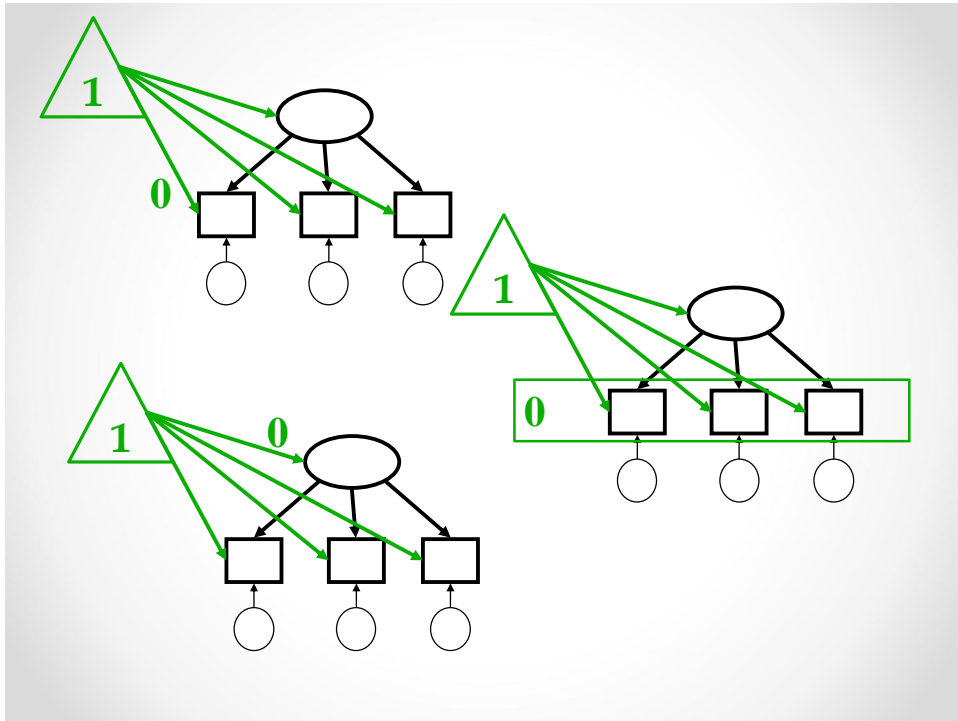
Setting the Scale?

Unless the units are meaningful, all methods are “almost” equivalent.

- **Referent Variable:** This scale is arbitrary and can change depending on the selected referent variable.
- **Latent Standardization:** The “natural” scale is lost, and the mean of the intercept factor becomes 0. This method also means that it is not the mean/variance of the first order factor that is constrained, but their intercepts/disturbances (as they are themselves explained by the second order intercept and slope factors). Another possibility is to fix the loading of a referent indicator to its value in the longitudinal CFA model, in order to freely estimate a variance that will be close to 1.
- **Effects Coding:** Ideal, but more prone to convergence problems.







Referent Indicator

Explicit

Model:

POSF BY q1 q2 q3 q4 q8;

NEGF BY q5 q6 q7;

[q1@0];

[q5@0];

[POSF*];

[NEGF*];

Means and Intercepts
are identified with
brackets []

Model:

POSF BY q1@1 q2 q3 q4 q8;

NEGF BY q5@1 q6 q7;

POSF*; NEGF*;

[q1@0];

[q5@0];

[POSF*];

[NEGF*];

Standardized factors

Model:

POSF BY q1* q2 q3 q4 q8;
NEGF BY q5* q6 q7;
POSF@1; NEGF@1;

By default, Mplus freely estimates all intercepts and constrain factors means to be 0.

Explicit

Model:

POSF BY q1* q2 q3 q4 q8;
NEGF BY q5* q6 q7;
POSF@1; NEGF@1;
[POSF@0];
[NEGF@0];

Effects Coding

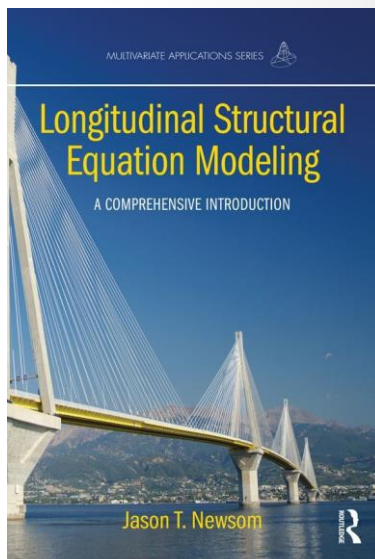
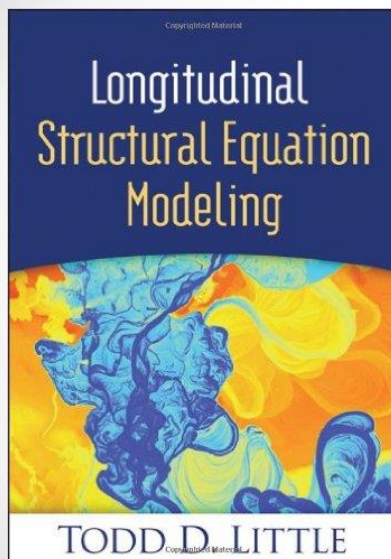
Model:

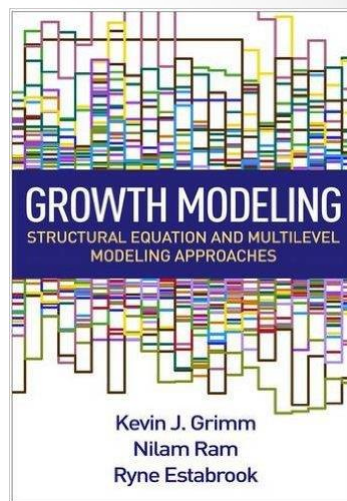
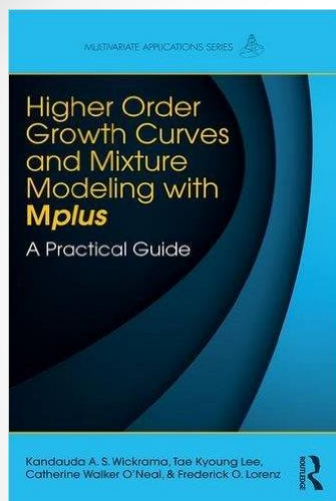
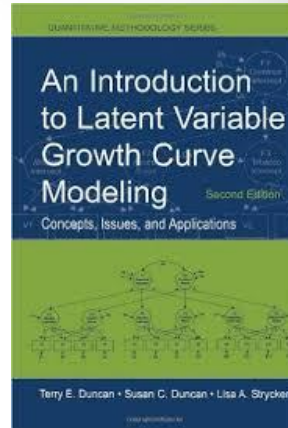
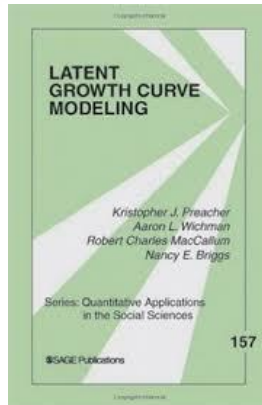
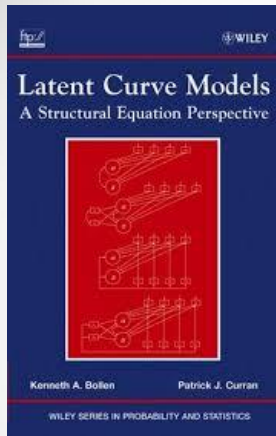
POSF BY q1* (11)
q2 q3 q4 q8 (12-15);
NEGF BY q5* (16)
q6 q7 (17-18);
POSF*; NEGF*;
[q1-q8] (i1-i8);
[POSF*]; [NEGF*];

MODEL CONSTRAINT:

11 = 5 - 12 - 13 - 14 - 15 ;
14 = 3 - 17 - 18;
i1 = 0-i2-i3-i4-i8;
i5 = 0-i6-i7;

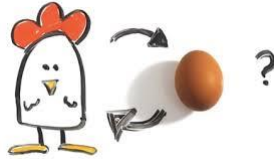
Longitudinal Structural Equation Modeling (continuous latent variables)





Why Longitudinal Data?

- Directionality



- Change



- Shape of change



Assumptions

- (1) The Meaning of Time.
- (2) Measurement Invariance.
- (3) Changes in Means and Variances.
- (4) Wording Effects (with Latent Models).
- (5) Predictive Equilibrium.

The Meaning(s) of Time

- **Temporal Ordering**: In some longitudinal studies, time simply serves to map temporal ordering: Year 1 of data collection, year 2, year 3, among a sample of participants of different age groups.
- **Meaningful Variable**: In some other studies, time is a meaningful variable and reflects the effect of age, tenure, etc.
- **Change**: Yet, in some other studies, time will reflect the impact of some natural (puberty, life transition) or experimental (intervention, experimentation, etc.) change that is also important.
- In an autoregressive cross-lagged models, time typically simply serves to map the temporal ordering of the relations. However, these analyses can also serve to test the effects of changes in a “pre-post” manner.
- When participants are really far apart in age, it might be useful to control for age in the estimation of the models.

The Meaning(s) of Time

- Over and above these considerations, many have noted the importance to think about the specific time interval that is used in a specific study, showing that it did importantly change the results.
- Not always possible.
- What do we know about the stability of the various constructs?
- Over which length of time can we expect change, both empirically and theoretically?
- How does that fit the research question.

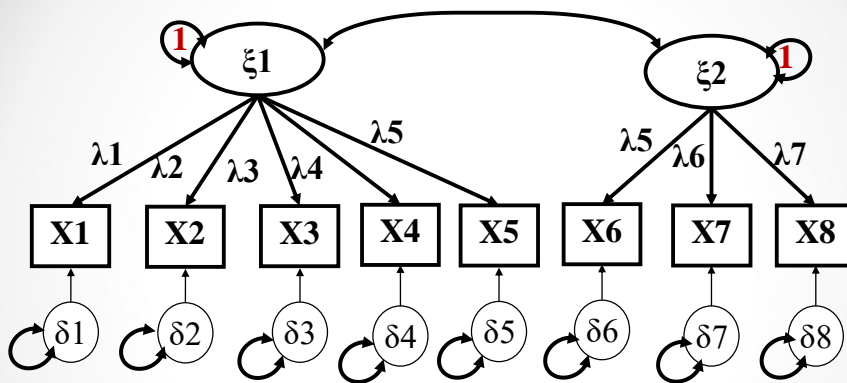
Cole, D.A., & Maxwell, S. E. (2003). Testing mediational models with longitudinal data: Questions and tips in the use of structural equation modeling. *Journal of Abnormal Psychology*, 112, 558–577.

Maxwell, S. E., & Cole, D.A. (2007). Bias in cross-sectional analyses of longitudinal mediation. *Psychological Methods*, 12, 23–44.

Selig, J.P. & Preacher, K.J. (2009). Mediation models for longitudinal data in developmental research. *Research in Human Development*, 6, 144–164.

Measurement invariance

- It is important to ascertain that the meaning of the constructs of interest has not changed over **time**.
- When **latent** constructs are used, and the only objective is to test for the **relations** among constructs, only weak (or partial weak) invariance (i.e. factor loadings) is required with continuous indicators.
- When **latent** constructs are used, and one also wants to test for latent **mean** differences, then strong (or partial strong) invariance (i.e. factors loadings and intercepts) is required.
- When **manifest scores** are used, then strict invariance (or partial strict (i.e., factors loadings, intercepts, and uniquenesses) is required.



Changes in Means and Variances

- A longitudinal measurement model can also be used to test for latent variances, latent covariances, and latent means differences as a function of time.
- This is the latent variable approach to repeated measures ANOVAs.
- When moving from a measurement model to a predictive model, even though only weak/strong invariance is required in a fully latent model, it remains useful to start from a model of strict invariance, and even from a model in which the latent variances and means are set to invariance when this is possible, as this induces parsimony in the model (maximising power).
- It also makes the variances and means easier to interpret (in SD units). Otherwise, they will still be interpretable in SD units of the first time point (if relying on the standardized factor approach).
- Not the covariances, as these do not have the same meaning in measurement versus predictive (residual covariances) models.

Changes in Means and Variances

- Sometimes it is not possible to rely on a fully latent predictive model: The complexity of the model becomes too great, and the model crashes. However, sometimes, it is still possible to estimate the fully latent measurement model.
- In these situations, invariance can, and should be tested.
- When this does not work, it should still be tested across pairs of time points.
- When invariance can be tested, it can be useful to rely on factor scores, rather than scale scores, as these provide a partial control for measurement errors, and can be saved from a model of strict invariance, or even of latent variances, covariances, and means invariance.

Wording Effects (Latent Models)

- In a longitudinal study, the same set of factor indicators are utilized at the different time points.
- This creates a methodological artefact given that what is “unique” to each item (uniqueness) becomes shared with what is unique to the same item measured at the other time points.
- Latent longitudinal models thus need to include correlated uniquenesses among matching indicators.
- Not doing so would risk inflating the autoregressive (stability) parameters.

Fully Latent Models

- In fully latent model, the decision to rely on the referent indicator, latent standardisation, or effects coding method will have a direct effect on the estimates of the average intercept ($[i]$), but also on the estimate of the means and variances of the other growth factors ($[s]$ will be expressed as deviation from $[i]$ in SD units if using the standardized factor approach).
- Preliminary: Longitudinal invariance of the repeated measures:
 - Requirement: Strong invariance.
 - Better: Strict invariance (makes the model more parsimonious)
 - Latent variance-covariances: not necessary, and results in invariant time-specific residuals in the LCM, not variances.
 - Latent means: invariance suggest that, on the average, there is no growth.
 - The fully latent model should start from a model of strong or strict invariance – maximum.

Longitudinal Invariance

FX_T1 BY X1_t1* X2_t1 X3_t1; FY_T1 BY Y1_t1* Y2_t1 Y3_t1; [X1_t1 X2_t1 X3_t1]; [Y1_t1 Y2_t1 Y3_t1]; X1_t1 X2_t1 X3_t1; Y1_t1 Y2_t1 Y3_t1; FX_T1@1; FY_T1@1; [FX_T1@0]; [FY_T1@0];	FX_T2 BY X1_t2* X2_t2 X3_t2; FY_T2 BY Y1_t2* Y2_t2 Y3_t2; [X1_t2 X2_t2 X3_t2]; [Y1_t2 Y2_t2 Y3_t2]; X1_t2 X2_t2 X3_t2; Y1_t2 Y2_t2 Y3_t2; FX_T2@1; FY_T2@1; [FX_T2@0]; [FY_T2@0];
--	--

X1_t1 X2_t1 X3_t1 **pwith** X1_t2 X2_t2 X3_t2;
! Equivalent to
! X1_t1 WITH X1_t2; X2_t1 WITH X2_t2; X3_t1 WITH X3_t2;

Weak Invariance

FX_T1 BY X1_t1* (I1)

X2_t1 X3_t1 (I2-I3);

FY_T1 BY Y1_t1* (I4)

Y2_t1 Y3_t1 (I5-I6);

[X1_t1 X2_t1 X3_t1];

[Y1_t1 Y2_t1 Y3_t1];

X1_t1 X2_t1 X3_t1;

Y1_t1 Y2_t1 Y3_t1;

FX_T1@I; FY_T1@I;

[FX_T1@0]; [FY_T1@0];

FX_T2 BY X1_t2* (I1)

X2_t2 X3_t2 (I2-I3);

FY_T2 BY Y1_t2* (I4)

Y2_t2 Y3_t2 (I5-I6);

[X1_t2 X2_t2 X3_t2];

[Y1_t2 Y2_t2 Y3_t2];

X1_t2 X2_t2 X3_t2;

Y1_t2 Y2_t2 Y3_t2;

FX_T2*; FY_T2*;

[FX_T2@0]; [FY_T2@0];

X1_t1 X2_t1 X3_t1 pwith X1_t2 X2_t2 X3_t2;

Strong Invariance

FX_T1 BY X1_t1* (I1)

X2_t1 X3_t1 (I2-I3);

FY_T1 BY Y1_t1* (I4)

Y2_t1 Y3_t1 (I5-I6);

[X1_t1 X2_t1 X3_t1] (i1-i3);

[Y1_t1 Y2_t1 Y3_t1] (i4-i6);

X1_t1 X2_t1 X3_t1;

Y1_t1 Y2_t1 Y3_t1;

FX_T1@I; FY_T1@I;

[FX_T1@0]; [FY_T1@0];

FX_T2 BY X1_t2* (I1)

X2_t2 X3_t2 (I2-I3);

FY_T2 BY Y1_t2* (I4)

Y2_t2 Y3_t2 (I5-I6);

[X1_t2 X2_t2 X3_t2] (i1-i3);

[Y1_t2 Y2_t2 Y3_t2] (i4-i6);

X1_t2 X2_t2 X3_t2;

Y1_t2 Y2_t2 Y3_t2;

FX_T2*; FY_T2*;

[FX_T2*]; [FY_T2*];

X1_t1 X2_t1 X3_t1 pwith X1_t2 X2_t2 X3_t2;

Strict Invariance

FX_T1 BY X1_t1* (I1)

X2_t1 X3_t1 (I2-I3);

FY_T1 BY Y1_t1* (I4)

Y2_t1 Y3_t1 (I5-I6);

[X1_t1 X2_t1 X3_t1] (i1-i3);

[Y1_t1 Y2_t1 Y3_t1] (i4-i6);

X1_t1 X2_t1 X3_t1 (u1-u3);

Y1_t1 Y2_t1 Y3_t1 (u4-u6);

FX_T1@I; FY_T1@I;

[FX_T1@0]; [FY_T1@0];

FX_T2 BY X1_t2* (I1)

X2_t2 X3_t2 (I2-I3);

FY_T2 BY Y1_t2* (I4)

Y2_t2 Y3_t2 (I5-I6);

[X1_t2 X2_t2 X3_t2] (i1-i3);

[Y1_t2 Y2_t2 Y3_t2] (i4-i6);

X1_t2 X2_t2 X3_t2 (u1-u3);

Y1_t2 Y2_t2 Y3_t2 (u4-u6);

FX_T2*; FY_T2*;

[FX_T2*]; [FY_T2*];

X1_t1 X2_t1 X3_t1 pwith X1_t2 X2_t2 X3_t2;

Var-Covar Invariance

FX_T1 BY X1_t1* (I1)

X2_t1 X3_t1 (I2-I3);

FY_T1 BY Y1_t1* (I4)

Y2_t1 Y3_t1 (I5-I6);

[X1_t1 X2_t1 X3_t1] (i1-i3);

[Y1_t1 Y2_t1 Y3_t1] (i4-i6);

X1_t1 X2_t1 X3_t1 (u1-u3);

Y1_t1 Y2_t1 Y3_t1 (u4-u6);

FX_T1@I; FY_T1@I;

FX_T1 WITH FY_T1 (c1);

[FX_T1@0]; [FY_T1@0];

FX_T2 BY X1_t2* (I1)

X2_t2 X3_t2 (I2-I3);

FY_T2 BY Y1_t2* (I4)

Y2_t2 Y3_t2 (I5-I6);

[X1_t2 X2_t2 X3_t2] (i1-i3);

[Y1_t2 Y2_t2 Y3_t2] (i4-i6);

X1_t2 X2_t2 X3_t2 (u1-u3);

Y1_t2 Y2_t2 Y3_t2 (u4-u6);

FX_T2@I; FY_T2@I;

FX_T2 WITH FY_T2 (c1);

[FX_T2*]; [FY_T2*];

X1_t1 X2_t1 X3_t1 pwith X1_t2 X2_t2 X3_t2;

Latent Means Invariance

FX_T1 BY X1_t1* (I1)
 X2_t1 X3_t1 (I2-I3);
 FY_T1 BY Y1_t1* (I4)
 Y2_t1 Y3_t1 (I5-I6);
 [X1_t1 X2_t1 X3_t1] (i1-i3);
 [Y1_t1 Y2_t1 Y3_t1] (i4-i6);
 X1_t1 X2_t1 X3_t1 (u1-u3);
 Y1_t1 Y2_t1 Y3_t1 (u4-u6);
 FX_T1@I; FY_T1@I;
 FX_T1 WITH FY_T1 (c1);
 [FX_T1@0]; [FY_T1@0];

FX_T2 BY X1_t2* (I1)
 X2_t2 X3_t2 (I2-I3);
 FY_T2 BY Y1_t2* (I4)
 Y2_t2 Y3_t2 (I5-I6);
 [X1_t2 X2_t2 X3_t2] (i1-i3);
 [Y1_t2 Y2_t2 Y3_t2] (i4-i6);
 X1_t2 X2_t2 X3_t2 (u1-u3);
 Y1_t2 Y2_t2 Y3_t2 (u4-u6);
 FX_T2@I; FY_T2@I;
 FX_T2 WITH FY_T2 (c1);
 [FX_T2@0]; [FY_T2@0];

X1_t1 X2_t1 X3_t1 pwith X1_t2 X2_t2 X3_t2;

Basic Longitudinal Models

Longitudinal Models.

When different variables are measured at each time point.

- Can be analyzed as cross-sectional data, with no added source of complexity, except in the treatment of missing data.
- This provides a way to analyse data while taking into account the temporal ordering of the data, but precluding clear tests of the directionality of the associations.
- **Basic rule:** Use all available cases, unless there is a specific reason not to (i.e., a key variable is completely missing, ethics/consents, etc.).

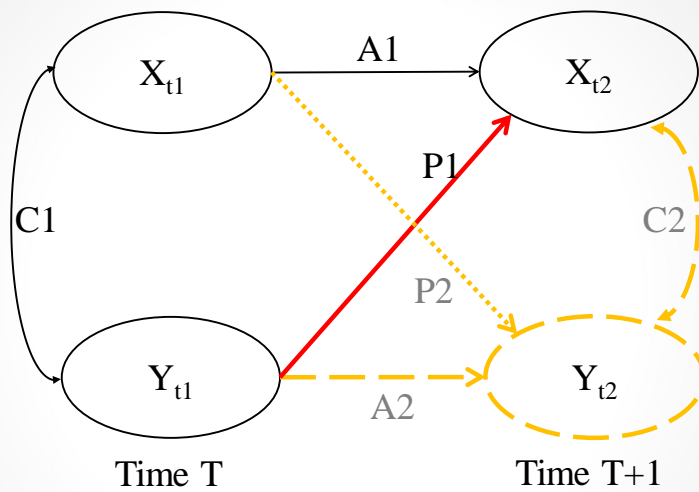
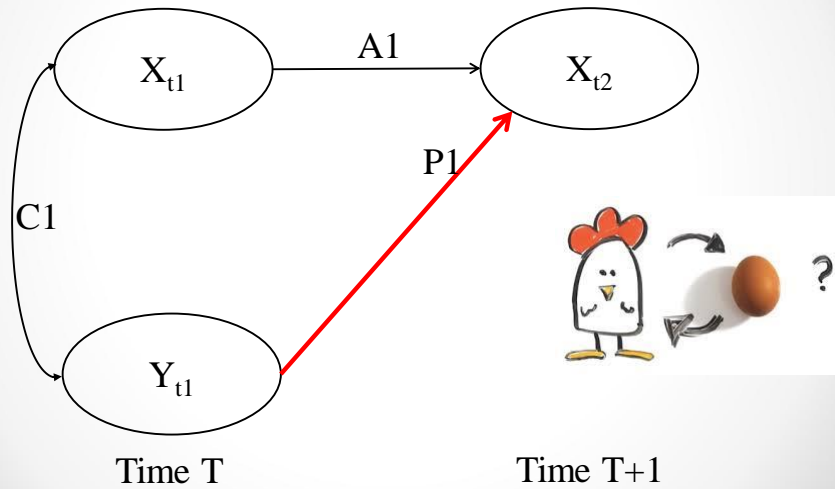
Longitudinal Models.

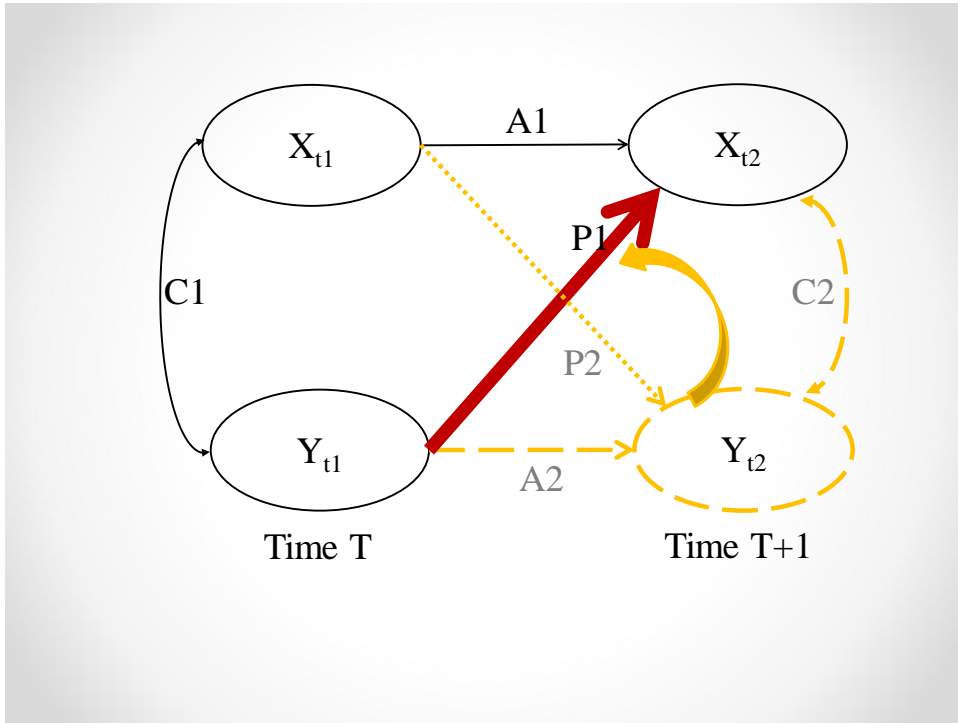
When the variables are measured at all time points.

Provides a way to control for baseline levels of the key variables of interest, and thus to model the “impact” of (or relation between) one variable and “changes” over time in the levels of an outcome variable (i.e., the influence of one variable on the part of the outcome that remains unexplained by baseline levels of that same outcome).

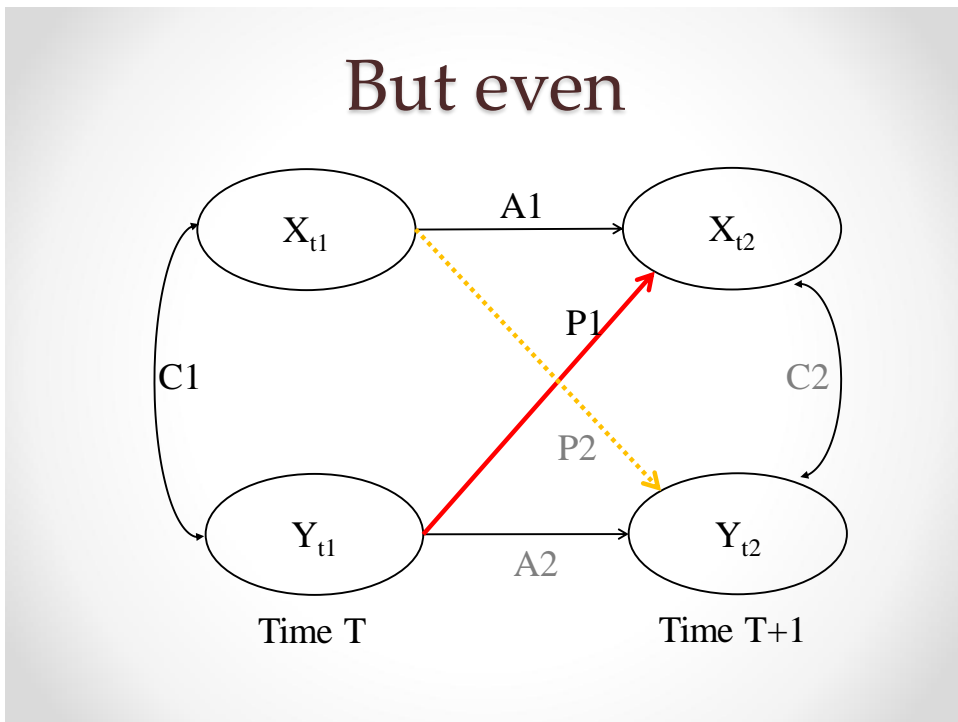
i.e., effects “over and above” the stability of the outcome.

The Classical Longitudinal Model

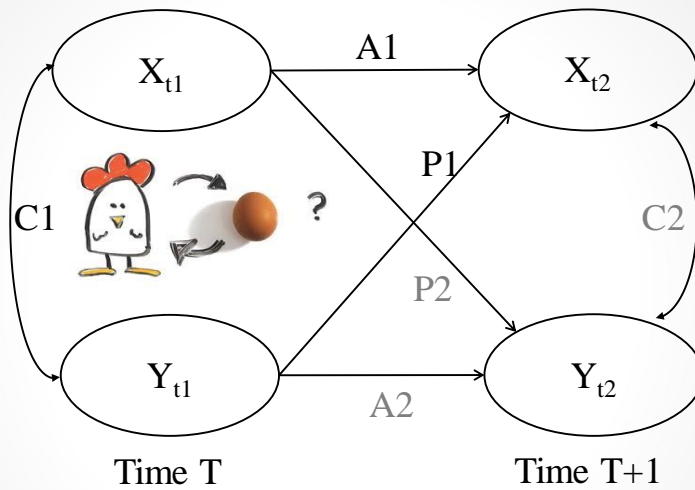
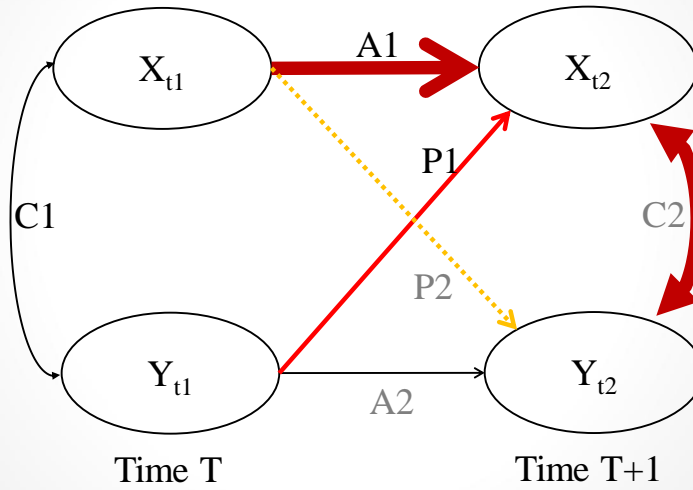




But even

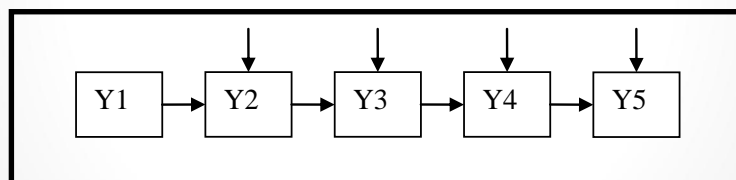


But even

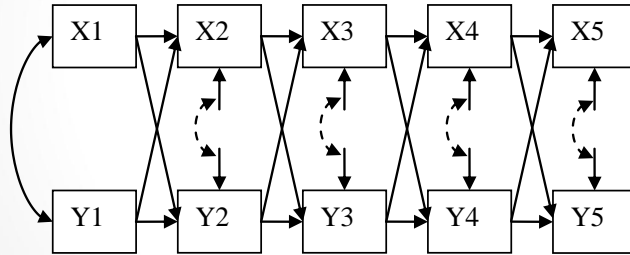


Autoregressive Cross Lagged Model

Autoregressive models allow one to estimate the stability / unstability of autoregressive parameters between adjacent measurement points over time.



Autoregressive cross-lagged models allow one to estimate the stability / unstability of autoregressive parameters between adjacent measurement points over time. They model the influence of “states” on later “states” but fail to consider stable process (traits).



$$y_{it} = \alpha_{y_t} + \rho_{y_t, y_{t-1}} y_{i,t-1} + \rho_{y_t, w_{t-1}} w_{i,t-1} + \varepsilon_{yit}$$

$$w_{it} = \alpha_{w_t} + \rho_{w_t, y_{t-1}} y_{i,t-1} + \rho_{w_t, w_{t-1}} w_{i,t-1} + \varepsilon_{wit}$$

$$y_{it} = \alpha_{yt} + \rho_{y_t, y_{t-1}} y_{i,t-1} + \rho_{y_t, w_{t-1}} w_{i,t-1} + \varepsilon_{yit}$$

$$w_{it} = \alpha_{wt} + \rho_{w_t, y_{t-1}} y_{i,t-1} + \rho_{w_t, w_{t-1}} w_{i,t-1} + \varepsilon_{wit}$$

Each process is thus a function of:

- An intercept (average level at Time 1) +

$$y_{it} = \alpha_{yt} + \rho_{y_t, y_{t-1}} y_{i,t-1} + \rho_{y_t, w_{t-1}} w_{i,t-1} + \varepsilon_{yit}$$

$$w_{it} = \alpha_{wt} + \rho_{w_t, y_{t-1}} y_{i,t-1} + \rho_{w_t, w_{t-1}} w_{i,t-1} + \varepsilon_{wit}$$

Each process is thus a function of:

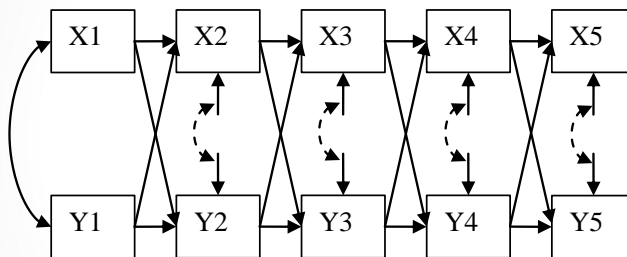
- An intercept (average level at Time 1) +
- A product of the autoregressive parameter and the previous time point of the variable itself +

$$y_{it} = \alpha_{y_t} + \rho_{y_t, y_{t-1}} y_{i,t-1} + \rho_{y_t, w_{t-1}} w_{i,t-1} + \varepsilon_{y_{it}}$$

$$w_{it} = \alpha_{w_t} + \rho_{w_t, y_{t-1}} y_{i,t-1} + \rho_{w_t, w_{t-1}} w_{i,t-1} + \varepsilon_{w_{it}}$$

Each process is thus a function of:

- **An intercept (average level at Time 1) +**
- **A product of the autoregressive parameter and the previous time point of the variable itself +**
- **A product of the cross lagged parameter and the previous time point of the other variable.**



TITLE: Non latent autoregressive cross lagged model, no constraints.

MODEL:

! Autoregressive part

X1;

X2 ON X1;

X3 ON X2;

X4 ON X3;

X5 ON X4;

Y1;Y2 ON Y1;Y3 ON Y2;Y4 ON Y3;Y5 ON Y4;

! Cross lagged part

X2 ON Y1; X3 ON Y2; X4 ON Y3; X5 ON Y4;

Y2 ON X1;Y3 ON X2;Y4 ON X3;Y5 ON X4;

! Time-specific correlations

X2 WITH Y2; X3 WITH Y3; X4 WITH Y4; X5 WITH Y5;

Predictive Equilibrium

- Cole and Maxwell (2003) underscored the importance of tests of predictive equilibrium, which consist of testing the extent to which the observed relations generalize to the various time intervals.
- A predictive system that has reached equilibrium shows relations that can be expected to generalize – i.e. the developmental process under study has stabilized for the period under study.
- Without these tests, it is impossible to determine whether variations in results reflect random sampling variations or meaningful developmental changes.
- A model that has reached equilibrium is more parsimonious, and thus more stable and has more power.

Predictive Equilibrium

Equilibrium of:

- Autoregressions
- Cross-lagged relations
- Time-specific correlations.

Cole, D.A., & Maxwell, S. E. (2003). Testing mediational models with longitudinal data: Questions and tips in the use of structural equation modeling. *Journal of Abnormal Psychology*, 112, 558–577.

Morin, A.J.S., Arens, A.K., Maiano, C., Ciarrochi, J., Tracey, D., Parker, P.D., & Craven, R.G. (In press Accepted, 12 September 2016). Reciprocal Relationships between Teacher Ratings of Internalizing and Externalizing Behaviors in Adolescents with Different Levels of Cognitive Abilities. *Journal of Youth and Adolescence*. Early view doi:10.1007/s10964-016-0574-3

MODEL:

! Autoregressive part

X1;

X2 ON X1 (a1); X3 ON X2 (a1); X4 ON X3 (a1); X5 ON X4 (a1);

Y1;

Y2 ON Y1 (a2); Y3 ON Y2 (a2); Y4 ON Y3 (a2); Y5 ON Y4 (a2);

! Cross lagged part

X2 ON Y1; X3 ON Y2; X4 ON Y3; X5 ON Y4;

Y2 ON X1; Y3 ON X2; Y4 ON X3; Y5 ON X4;

! Time-specific correlations

X2 WITH Y2; X3 WITH Y3; X4 WITH Y4; X5 WITH Y5;

MODEL:

! Autoregressive part

X1;

X2 ON X1 (a1); X3 ON X2 (a1); X4 ON X3 (a1); X5 ON X4 (a1);

Y1;

Y2 ON Y1 (a2); Y3 ON Y2 (a2); Y4 ON Y3 (a2); Y5 ON Y4 (a2);

! Cross lagged part

X2 ON Y1 (CL1); X3 ON Y2 (CL1); X4 ON Y3 (CL1); X5 ON Y4 (CL1);

Y2 ON X1 (CL2); Y3 ON X2 (CL2); Y4 ON X3 (CL2); Y5 ON X4 (CL2);

! Time-specific correlations

X2 WITH Y2; X3 WITH Y3; X4 WITH Y4; X5 WITH Y5;

MODEL:

! Autoregressive part

X1;

X2 ON X1 (a1); X3 ON X2 (a1); X4 ON X3 (a1); X5 ON X4 (a1);

Y1;

Y2 ON Y1 (a2); Y3 ON Y2 (a2); Y4 ON Y3 (a2); Y5 ON Y4 (a2);

! Cross lagged part

X2 ON Y1 (CL1); X3 ON Y2 (CL1); X4 ON Y3 (CL1); X5 ON Y4 (CL1);

Y2 ON X1 (CL2); Y3 ON X2 (CL2); Y4 ON X3 (CL2); Y5 ON X4 (CL2);

! Time-specific correlations

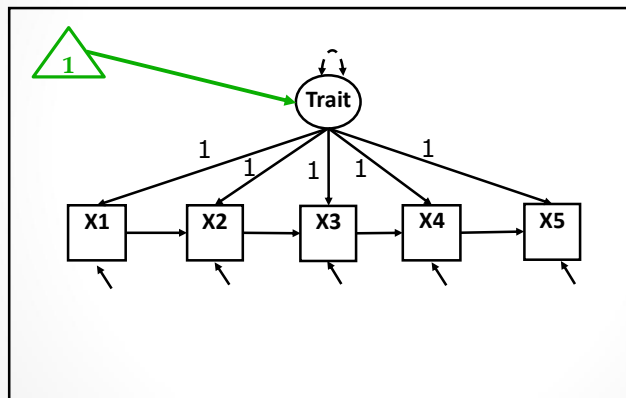
X2 WITH Y2 (CI); X3 WITH Y3 (CI);

X4 WITH Y4 (CI); X5 WITH Y5 (CI);

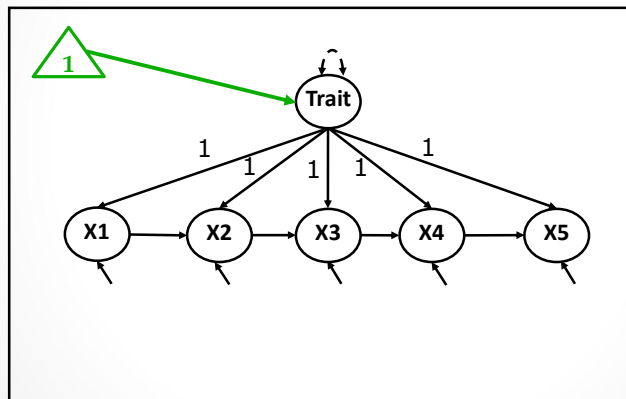
Limitations of the Autoregressive Cross Lagged Model

(1) Models the influence of *states* on later *states* but fail to consider stable developmental processes (*traits*).

State Trait Models

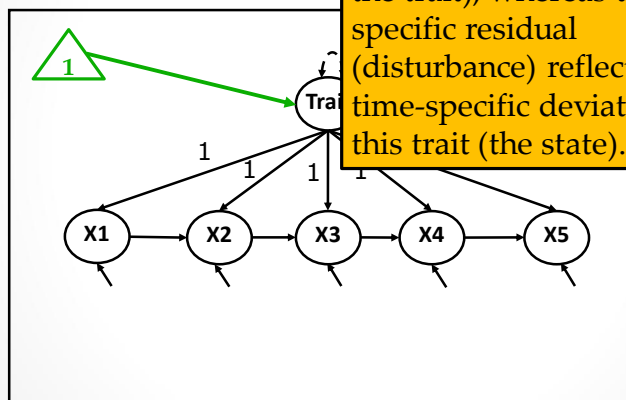


State-Trait-Error Models?



State-Trait-Error

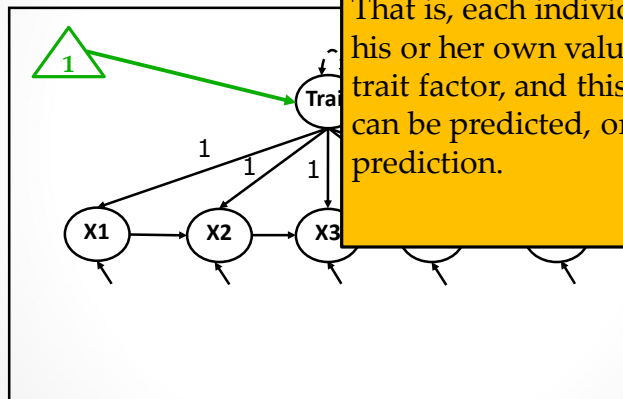
In these models the trait factor reflects the average level observed over time in the repeated measures (i.e., the trait), whereas the time specific residual (disturbance) reflects the time-specific deviation from this trait (the state).



State-Trait-Error

The trait factor is what is called a “random” effect.

That is, each individual has his or her own value on this trait factor, and this value can be predicted, or used in prediction.



MODEL:

!Trait Factor

Trait BY X1@1 X2@1 X3@1 X4@1 X5@1;

Trait*; [Trait*];

[X1@0 X2@0 X3@0 X4@0 X5@0];

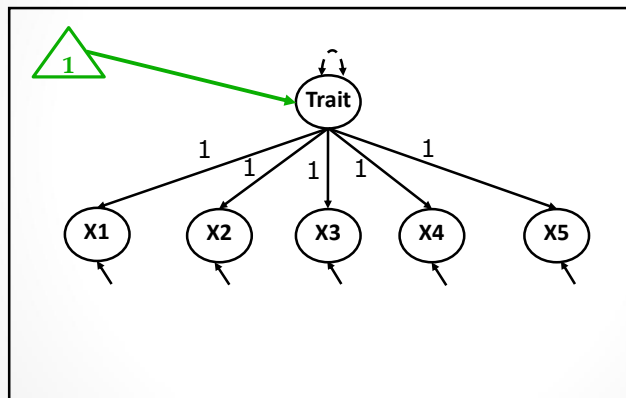
X1 X2 X3 X4 X5;

!Autoregressive part

X2 ON X1; X3 ON X2; X4 ON X3; X5 ON X4;

- Kenny, D.A., & Zautra, A. (1995). The trait-state-error model for multiwave data. *Journal of Consulting and Clinical Psychology*, 63, 52-59.
- Steyer, R., Schmitt, M., & Eid, M. (1999). Latent State-Trait theory and research in personality and individual differences. *European Journal of Personality*, 13, 389-408.
- Geiser, C., Keller, B.T., & Lockhart, G. (2013). First- versus second-order latent growth curve models Some insights from latent state-trait theory. *Structural Equation Modeling*, 20, 479-503.
- Cole, D.A., Martin, N.C., & Steiger, J.H. (2005). Empirical and conceptual problems with longitudinal trait-state- models: Introducing a trait-state-occasion model. *Psychological Methods*, 10, 3-20.
- Ciesla, J.A., Cole, D.A., & Steiger, J.H. (2007). Extending the trait-state-occasion model: How important is within-wave measurement equivalence. *Structural Equation Modeling*, 14, 77-97.

[Random Intercept Model]



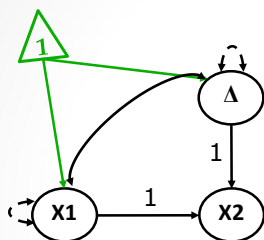
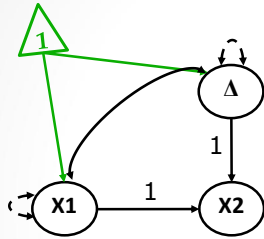
Limitations of the Autoregressive Cross-Lagged and State-Trait Models

- (2) Do not explicitly model developmental *change* over time.
- (3) State-Trait Models neglect to account for evolution occurring at the Trait level.

Latent Change Models

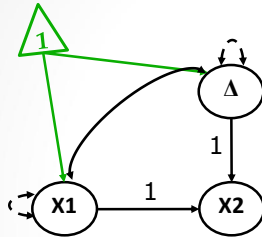
- Difference scores (or change scores) have been around for a long time in applied research.
- $\text{Change} = \text{Time 2 rating} - \text{Time 1 rating}$.
- But these models have come under major criticisms, due to the fact that change scores result in an multiplicative combination of measurement errors.
- Latent changes scores solve these problems.

McArdle, J.J. (2009). Latent Variable Modeling of Differences and Changes with Longitudinal Data. *Annual Review of Psychology*, 60, 577-605.



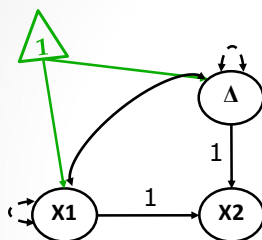
By fixing the regression of $X2$ on $X1$ to be exactly 1, you force the residual of $X2$ to correspond exactly to the amount of change since $T1$.

By forcing the loading of $X2$ on the change factor to be exactly 1 (as well as fixing the intercept and residual of $X2$ to be exactly 0), you force all information present in $X2$ to be absorbed in the change factor.

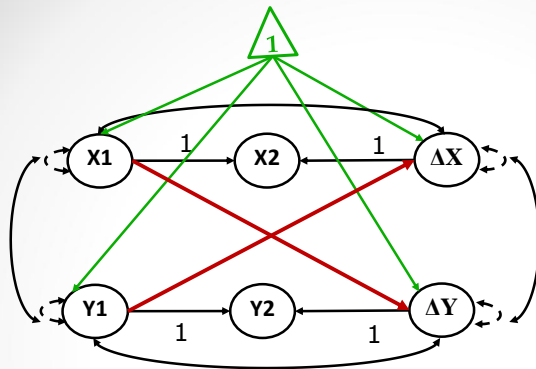
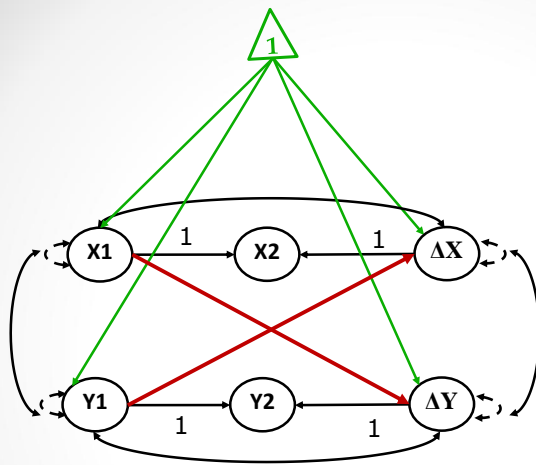


Without the change factor:

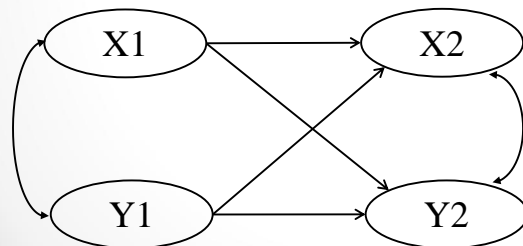
- (1) Other variables would still predict $X2$. With the change factor, they are forced to predict only $X2 - X1$.
- (2) It would be impossible to assess the correlation between the initial level ($X1$) and the change factor (reflecting the fact that the rate of change might be more or less pronounced depending on initial levels).



The change factor is random; i.e. it has a mean (reflecting the average change in the total sample) and a variance (reflecting inter-individual differences in rates of change).



Versus?



Illustration

- Using the longitudinal data set, focusing on Time 1 and Time 2 measures of loneliness and self-esteem, and starting with a model of strong invariance based on the referent indicator approach.
- **Autoregressive cross-lagged model:**
- **Latent change model:**

The latent part of both models
Referent indicator method

LON_t1 BY Lon1_t1@1

Lon2_t1 Lon3_t1 Lon4_t1 Lon5_t1 (12-15);

LON_T2 BY Lon1_t2@1

Lon2_t2 Lon3_t2 Lon4_t2 Lon5_t2 (12-15);

[Lon1_t1@0];

[Lon2_t1 Lon3_t1 Lon4_t1 Lon5_t1] (11-15);

[Lon1_t2@0];

[Lon2_t2 Lon3_t2 Lon4_t2 Lon5_t2] (11-15);

Lon1_t1 Lon2_t1 Lon3_t1 Lon4_t1 Lon5_t1;! Add (u1-u5) for strict invariance

Lon1_t2 Lon2_t2 Lon3_t2 Lon4_t2 Lon5_t2;! Add (u1-u5) for strict invariance

Lon1_t1 Lon2_t1 Lon3_t1 Lon4_t1 Lon5_t1 PWITH Lon1_t2
Lon2_t2 Lon3_t2 Lon4_t2 Lon5_t2;

```

se_t1 BY se_t1@1
se2_t1 se3_t1 se4_t1 se5_t1 (112-115);
se_T2 BY se_t2@1
se2_t2 se3_t2 se4_t2 se5_t2 (112-115);

```

```

[se_t1@0];
[se2_t1 se3_t1 se4_t1 se5_t1] (i12-i15);
[se_t2@0];
[se2_t2 se3_t2 se4_t2 se5_t2] (i12-i15);

```

```

se_t1 se2_t1 se3_t1 se4_t1 se5_t1; !Add (u11-u15) for strict invariance
se_t2 se2_t2 se3_t2 se4_t2 se5_t2; !Add (u11-u15) for strict invariance

```

```

se_t1 se2_t1 se3_t1 se4_t1 se5_t1 PWITH se_t2 se2_t2
se3_t2 se4_t2 se5_t2;

```

Autoregressive-Cross Lagged

```

LON_t1*;
LON_T2*;
[LON_t1*];
[LON_T2*];

```

```

se_t1*;
se_T2*;
[se_t1*];
[se_T2*];

```

```

se_T2 ON se_t1 LON_t1;
LON_T2 ON LON_t1 se_t1;
LON_t1 WITH se_t1;
LON_t2 WITH se_t2;

```

Latent Change Model

se_T2 ON se_t1@1;	[changepse*] ;
changepse BY se_T2@1;	[changelon*];
	changepse*;
LON_T2 ON LON_t1@1;	changelon*;
changelon BY LON_T2@1;	
	se_T2@0;
changepse WITH changelon se_t1 ;	LON_T2@0;
changelon WITH LON_t1;	[se_T2@0];
se_t1 WITH LON_t1;	[LON_T2@0];
changepse ON LON_t1;	se_T1*;
changelon ON se_t1;	LON_T1*;
	[se_T1*];
	[LON_T1*];

Illustration

- **Autoregressive cross-lagged model:**

$\chi^2 = 534.445$; $df = 170$; CFI = .935; TLI = .927; RMSEA = .038.

- **Latent change model:**

$\chi^2 = 534.445$; $df = 170$; CFI = .935; TLI = .927; RMSEA = .038

Illustration

- **Autoregressive cross-lagged model:**

$\chi^2 = 534.445$; $df = 170$; $CFI = .935$; $TLI = .927$; $RMSEA = .038$.

Self-Esteem (t1) \rightarrow Loneliness (t2): $\beta = -.045$

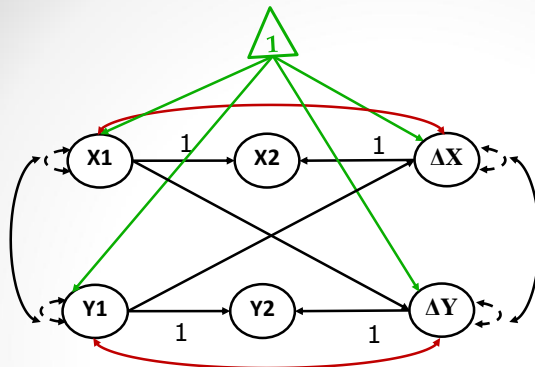
Loneliness (t1) \rightarrow Self-Esteem (t2): $\beta = .070^*$

- **Latent change model:**

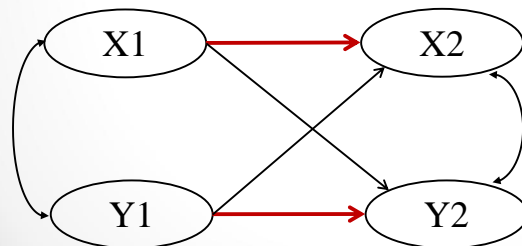
$\chi^2 = 534.445$; $df = 170$; $CFI = .935$; $TLI = .927$; $RMSEA = .038$

Self-Esteem (t1) $\rightarrow \Delta$ Loneliness (t1-t2): $\beta = .186^{**}$

Loneliness (t1) $\rightarrow \Delta$ Self-Esteem (t1-t2): $\beta = .179^{**}$



BUT ?

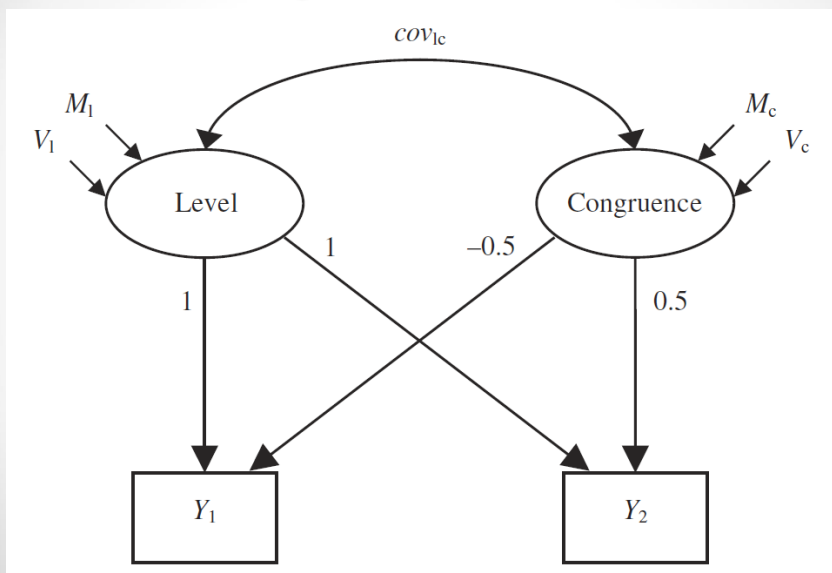


Parenthesis: Other Applications of Latent Change Scores with Cross Sectional Data

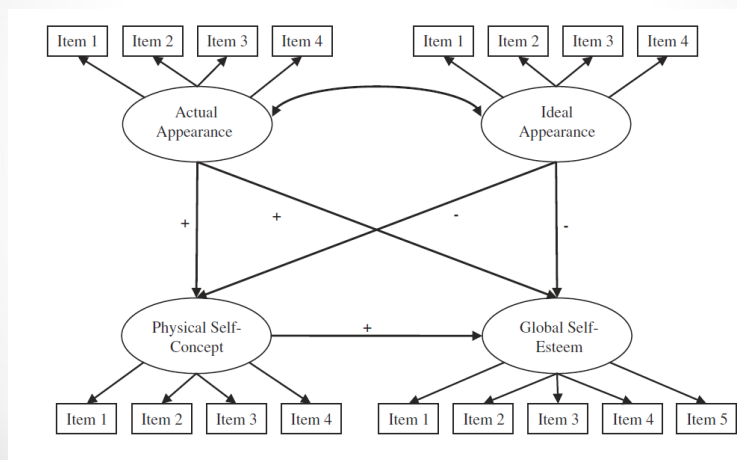
Utility

- Person-Environment fit
- Actual-Ideal Discrepancies
- Etc.
- Any domain when there is a reason to believe that discrepancies between one construct and another is a meaningful variable to consider

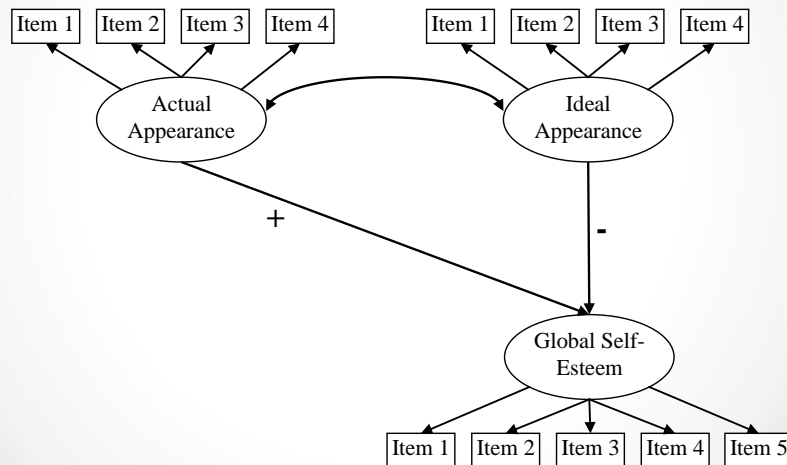
Latent Congruence (Cheung, 2009)



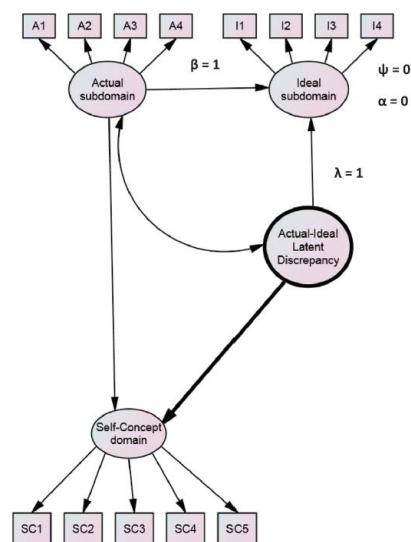
Actual-Ideal Discrepancies (Morin et al., 2015)



Actual-Ideal Discrepancies (Morin et al., 2015)



Or Latent Discrepancy Model (Scalas et al., 2014) ?



Morin et al. (2015)

- **Classical Actual-Ideal Discrepancy Model**

$\chi^2 = 789.21$, $df = 113$; CFI = 0.94; TLI = 0.93; RMSEA = 0.06

Ideal \rightarrow PSC: $b = -0.079$; $p \leq .01$

Ideal \rightarrow GSE: $b = -0.006$, ns

- **Latent Discrepancy Model**

$\chi^2 = 789.21$, $df = 113$; CFI = 0.94; TLI = 0.93; RMSEA = 0.06

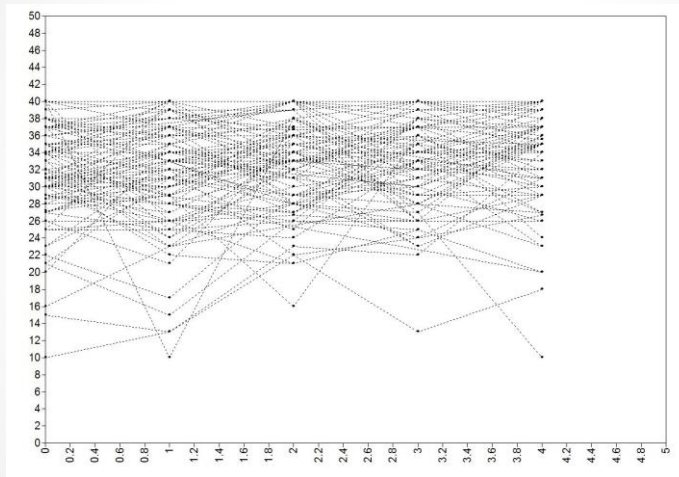
Discrepancy \rightarrow PSC: $b = -0.079$; $p \leq .01$

Discrepancy \rightarrow GSE: $b = -0.006$, ns

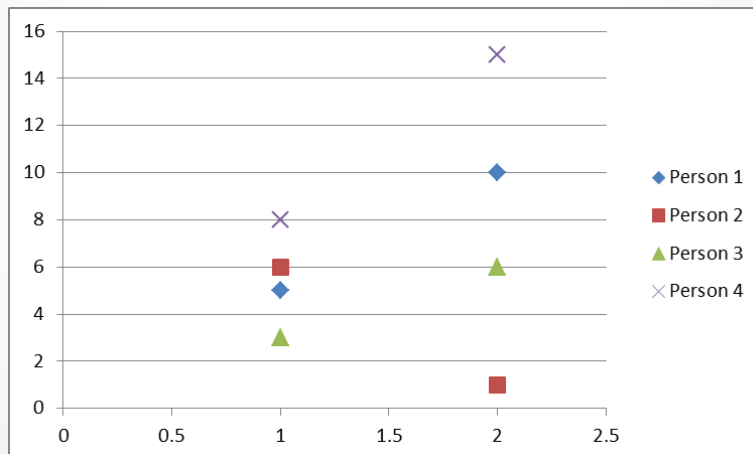
Limitations of the Latent Change Models

(4) Are unable to depict the shape of the growth trajectory over time. They do model “change”, but not the process of change

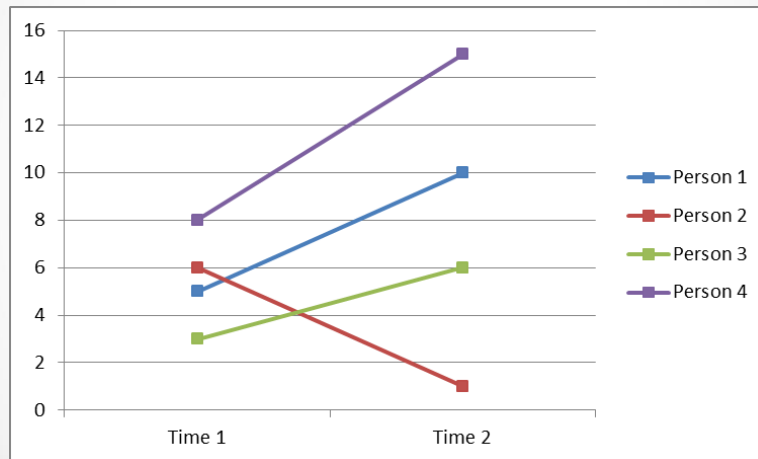
Individual trajectories



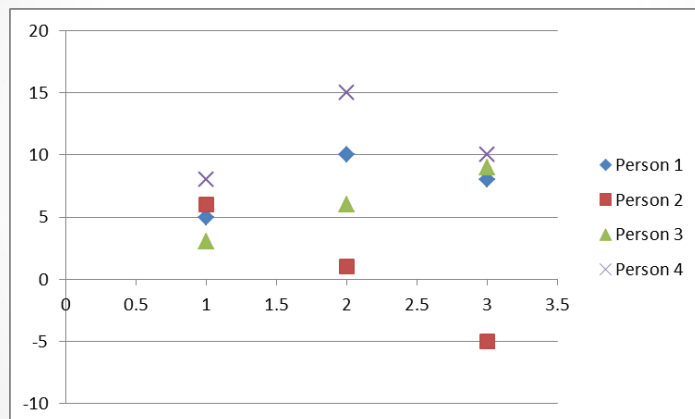
Two time points



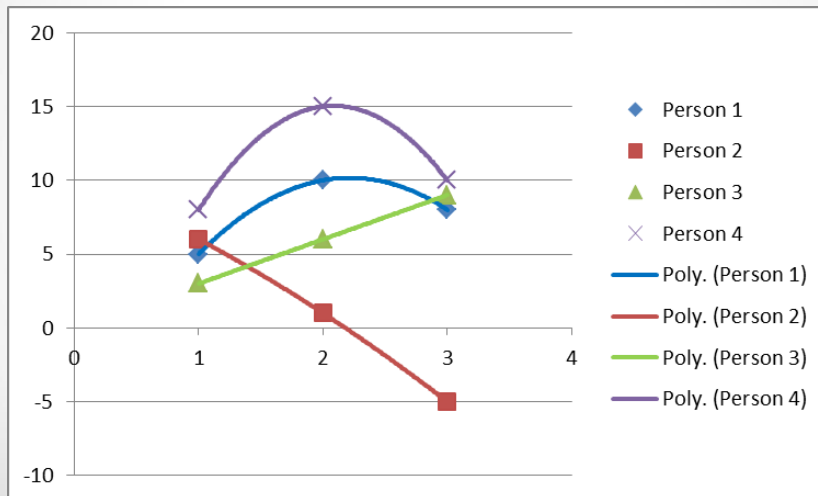
Two time points



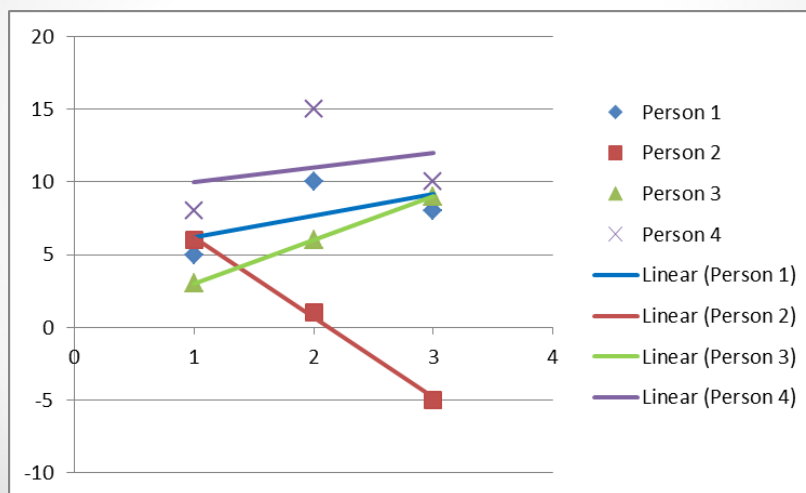
Three Time points



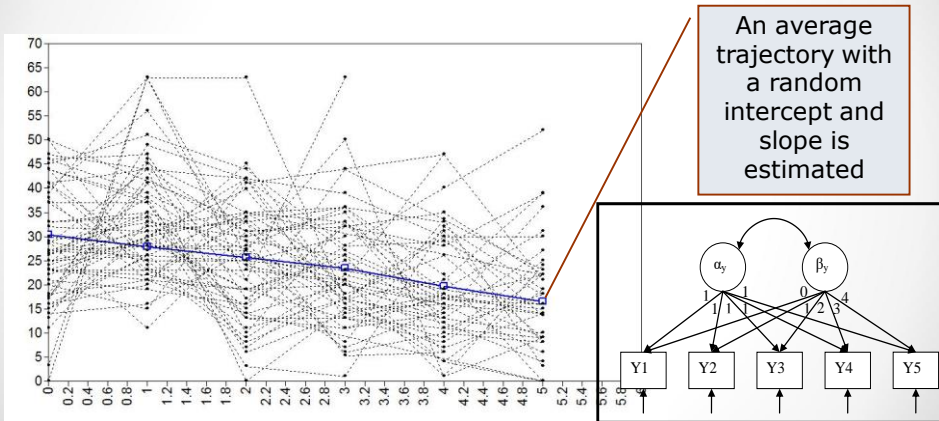
Three Time points



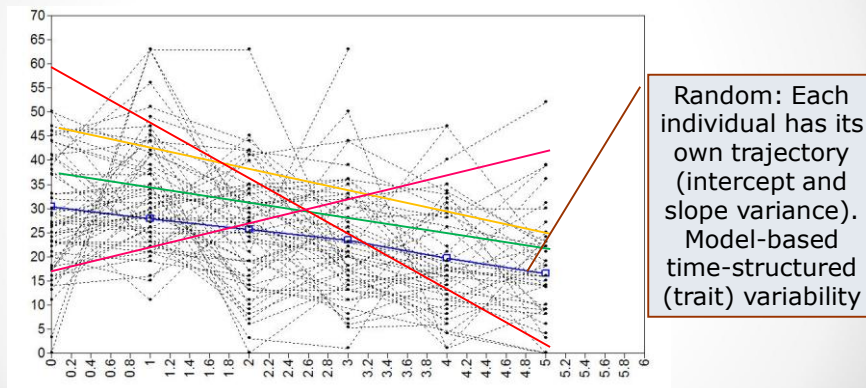
Three Time points



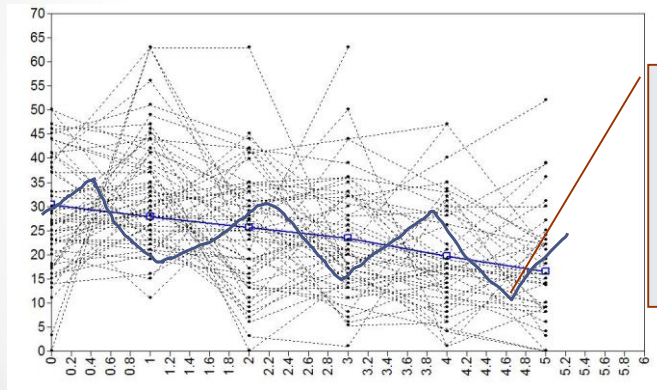
Latent curve models allow one to synthesize individual trajectories with few latent parameters through a restricted factor model.



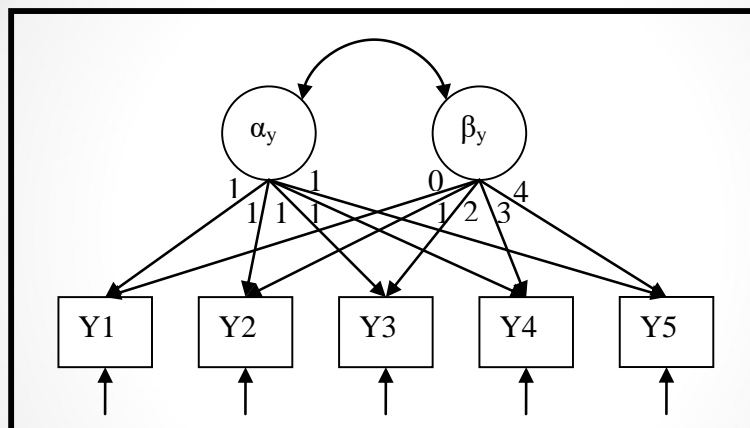
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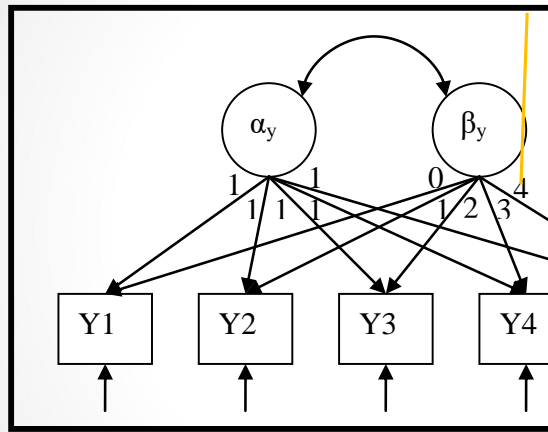


Latent curve models allow one to synthesize individual trajectories with few latent parameters through a restricted factor model.



Residuals:
Few persons
follow a perfectly
linear
(curvilinear, etc.)
trajectory.
Residual state-
like variability.





Time codes (λ): Reflects the passage of time.

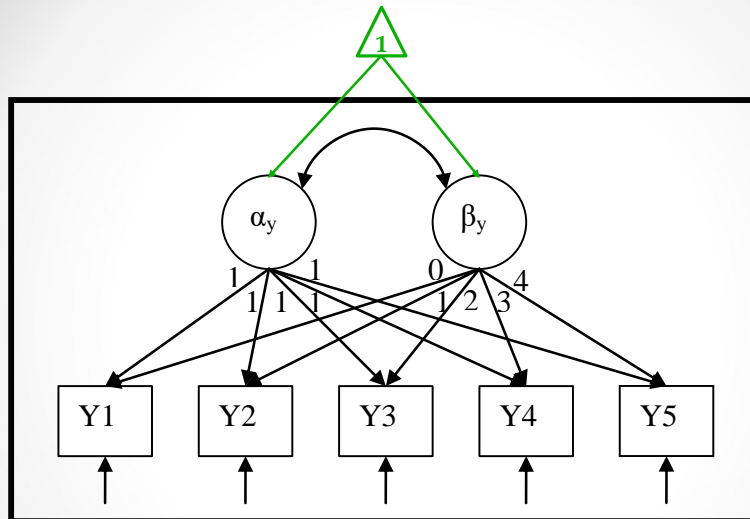
Linear: 0-1-2-3-4-5 (or some other reflecting the true passage of time, e.g. 0-1-3-4-6).

Quadratic: Addition of a second slope factor with squared time codes (0-1-4-9-16-25).

Piecewise: Two slopes with loadings reflecting the transition point:

- 0-1-2-3-3-3-3
- 0-0-0-0-1-2-3

Multibase: Free loadings reflecting the % of change between 0 and 1.



Linear LCM

$$Y_{it} = \alpha_{iy} + \beta_{iy}\lambda_t + \varepsilon_{yit}$$

$$\alpha_{iy} = \mu_{\alpha} + \zeta_{\alpha yi}$$

$$\beta_{iy} = \mu_{\beta} + \zeta_{\beta yi}$$

Requires 3 time points

Slope factor

Time specific residuals

$$Y_{it} = \alpha_{iy} + \beta_{iy}\lambda_t + \varepsilon_{yit}$$

Intercept factor

Time codes (λ): Reflects the passage of time.

Linear: 0-1-2-3-4-5 (or some other reflecting the true passage of time, e.g. 0-1-3-4-6).

$$\alpha_{iy} = \mu_{\alpha} + \zeta_{\alpha yi}$$

$$\beta_{iy} = \mu_{\beta} + \zeta_{\beta yi}$$

Average intercept and slope
observed in the total sample

$$\alpha_{iy} = \mu_{\alpha} + \zeta_{\alpha yi}$$

$$\beta_{iy} = \mu_{\beta} + \zeta_{\beta yi}$$

Disturbances (zeta)
reflecting the variability of
the estimated intercepts and
slopes across cases within
latent trajectory classes.
These disturbances have a
mean of zero and a
variance/covariance matrix
represented by Phi:

$$\Phi_y = \begin{bmatrix} \psi_{\alpha\alpha y} & \\ \psi_{\alpha\beta y} & \psi_{\beta\beta y} \end{bmatrix}$$

Quadratic LCM

$$Y_{it} = \alpha_{iy} + \beta 1_{iy} \lambda_t + \beta 2_{iy} \lambda_t^2 + \varepsilon_{yit}$$

$$\alpha_{iy} = \mu_{\alpha} + \zeta_{\alpha yi}$$

$$\beta 1_{iy} = \mu_{\beta 1} + \zeta_{\beta 1 yi}$$

$$\beta 2_{iy} = \mu_{\beta 2} + \zeta_{\beta 2 yi}$$

Requires 4 time points

$$Y_{it} = \alpha_{iy} + \beta 1_{iy} \lambda_t + \beta 2_{iy} \lambda_t^2 + \varepsilon_{yit}$$

Time codes (λ): Reflects the passage of time.

Linear: 0-1-2-3-4-5 (or some other reflecting the true passage of time, e.g. 0-1-3-4-6).

Quadratic: Addition of a second slope factor with squared time codes (0-1-4-9-16-25).

$$\alpha_{iy} = \mu_{\alpha} + \zeta_{\alpha yi}$$

$$\beta 1_{iy} = \mu_{\beta 1} + \zeta_{\beta 1 yi}$$

$$\beta 2_{iy} = \mu_{\beta 2} + \zeta_{\beta 2 yi}$$

Disturbances (zeta) reflecting the variability of the estimated intercepts and slopes across cases within latent trajectory classes. These disturbances have a mean of zero and a variance/covariance matrix represented by Phi:

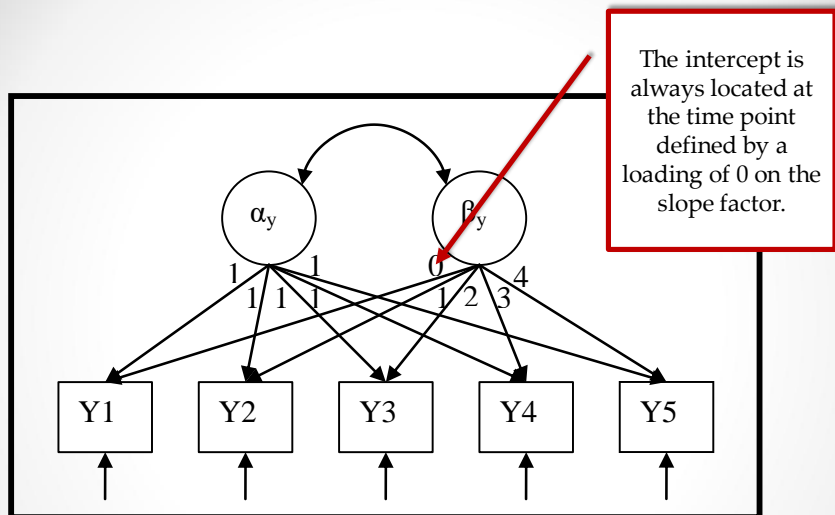
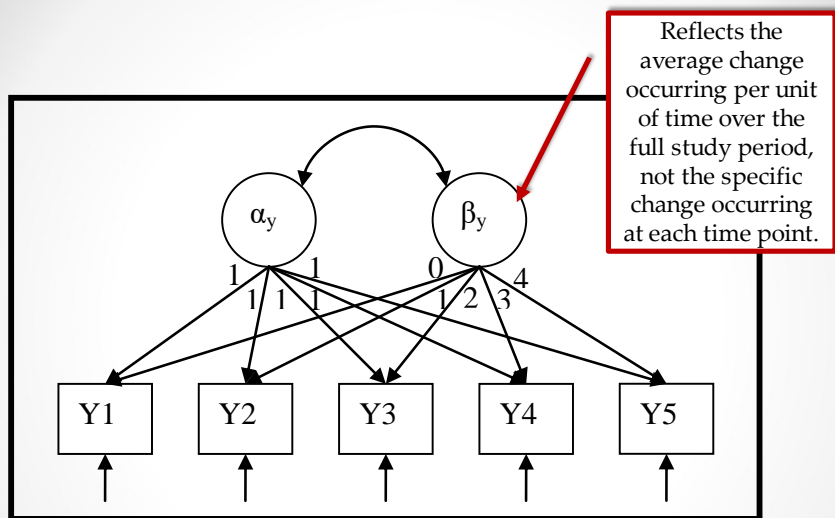
$$\Phi_y = \begin{bmatrix} \psi_{\alpha\alpha y} & & \\ \psi_{\alpha\beta 1y} & \psi_{\beta 1\beta 1y} & \\ \psi_{\alpha\beta 2y} & \psi_{\beta 1\beta 2y} & \psi_{\beta 2\beta 2y} \end{bmatrix}$$

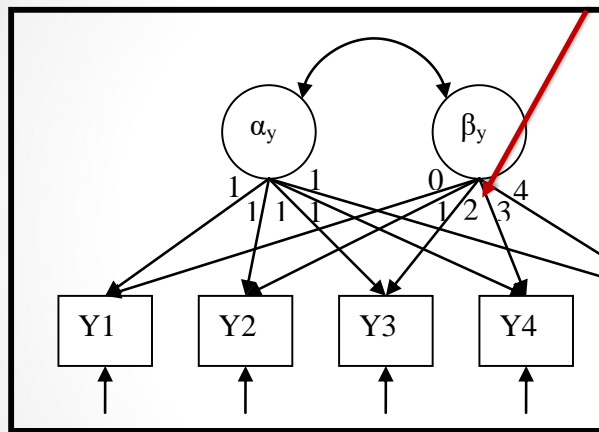
Time

- In any applications of LCM time must be a meaningful variable: Age, Grade, Tenure, etc.
- Simply having a sample of participants of various age/grade/tenure level measured repeatedly over time is not sufficient. One needs to expect that time plays a meaningful role in order to support the needs to rely on analyses of trajectories.
- The argument that LCM are the best way to assess change over time is fallacious, unless time is meaningful.
- The **position of the intercept** should never be arbitrary, given that it will have a great impact on the results.
- The **time codes** are also not arbitrary, and need to reflect the passage of time.

Mehta, P. D., & West, S. G. (2000). Putting the individual back into individual growth curves. *Psychological Methods*, 5, 23–43.

Biesanz, J. C., Deeb-Soosa, N., Papadakis, A. A., Bollen, K. A., & Curran, P. J. (2004). The role of coding time in estimating and interpreting growth curve models. *Psychological Methods*, 9, 30–52.





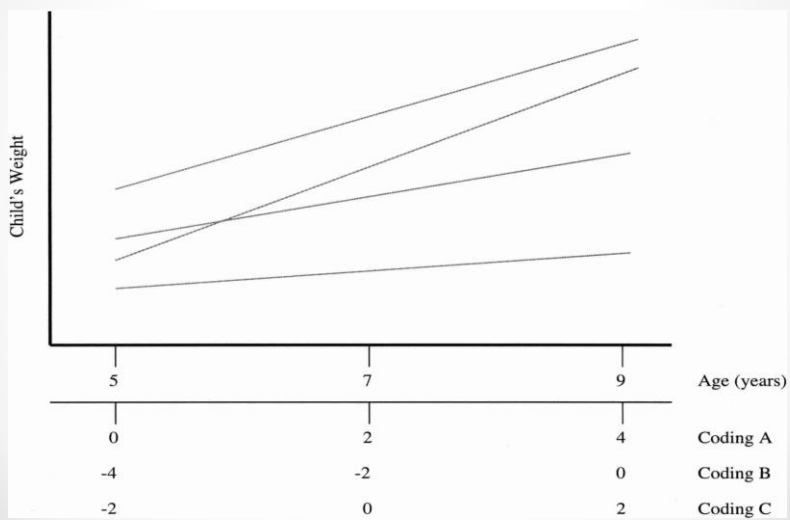
For example, if Time 1, 2, 3 are taken 6 months apart, and time 4 one year after time 3, and time 5 two years after time 4, then the time code should be:

0/.5/1/2/4

If time 1 reflects background control measures and the intercept should really be located at Time 2:

-.5/0/.5/1.5/3.5

Biesanz et al., (2004)

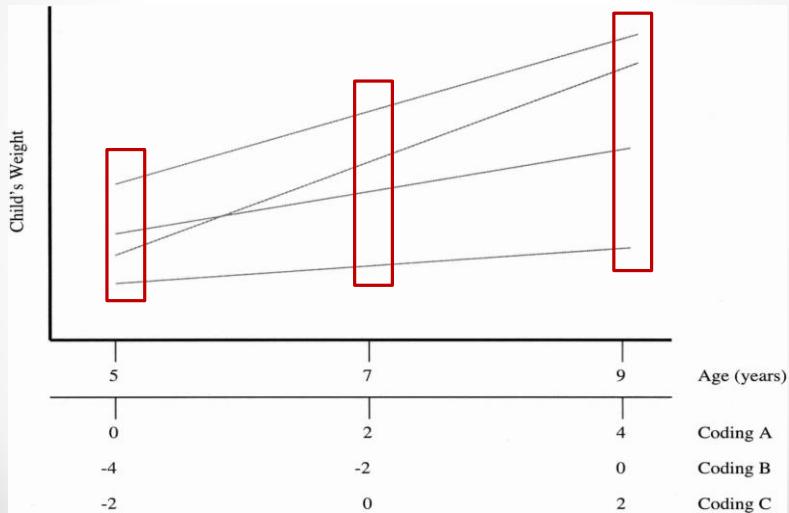


$$\Lambda_A = \begin{bmatrix} 1 & 0 \\ 1 & 2 \\ 1 & 4 \end{bmatrix}, \Lambda_B = \begin{bmatrix} 1 & -4 \\ 1 & -2 \\ 1 & 0 \end{bmatrix}, \Lambda_C = \begin{bmatrix} 1 & -2 \\ 1 & 0 \\ 1 & 2 \end{bmatrix}.$$

Parameter	Model					
	A		B		C	
	Estimate	SE	Estimate	SE	Estimate	SE
Variance						
Intercept	28.776	6.143	260.032	33.037	111.599	13.522
Linear	8.201	1.470	8.201	1.470	8.201	1.470
Covariance						
Intercept–linear	12.505	2.676	45.309	6.382	28.907	3.912
<i>M</i>						
Intercept	39.457	0.487	71.707	1.369	55.582	0.877
Linear	8.063	0.267	8.063	0.267	8.063	0.267
Unique variance						
Age 5	8.319	5.280	8.319	5.280	8.319	5.280
Age 7	12.094	4.880	12.094	4.880	12.094	4.880
Age 9	57.168	15.992	57.168	15.992	57.168	15.992
Model $T_{ML}(df = 1)$	2.131		2.131		2.131	

$$\Lambda_A = \begin{bmatrix} 1 & 0 \\ 1 & 2 \\ 1 & 4 \end{bmatrix}, \Lambda_B = \begin{bmatrix} 1 & -4 \\ 1 & -2 \\ 1 & 0 \end{bmatrix}, \Lambda_C = \begin{bmatrix} 1 & -2 \\ 1 & 0 \\ 1 & 2 \end{bmatrix}.$$

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Age 7	12.094	4.880	12.094	4.880	12.094	4.880
Age 9	57.168	15.992	57.168	15.992	57.168	15.992
Model $T_{ML}(df = 1)$	2.131		2.131		2.131	



$$\Lambda_A = \begin{bmatrix} 1 & 0 \\ 1 & 2 \\ 1 & 4 \end{bmatrix}, \Lambda_B = \begin{bmatrix} 1 & -4 \\ 1 & -2 \\ 1 & 0 \end{bmatrix}, \Lambda_C = \begin{bmatrix} 1 & -2 \\ 1 & 0 \\ 1 & 2 \end{bmatrix}.$$

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	A		B		C	
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M						
Intercept	39.457	0.487	71.707	1.369	55.582	0.877
Linear	8.063	0.267	8.063	0.267	8.063	0.267
Unique variance						
Age 5	8.319	5.280	8.319	5.280	8.319	5.280
Age 7	12.094	4.880	12.094	4.880	12.094	4.880
Age 9	57.168	15.992	57.168	15.992	57.168	15.992
Model $T_{ML}(df = 1)$	2.131		2.131		2.131	

$$\Lambda_A = \begin{bmatrix} 1 & 0 \\ 1 & 2 \\ 1 & 4 \end{bmatrix}, \Lambda_B = \begin{bmatrix} 1 & -4 \\ 1 & -2 \\ 1 & 0 \end{bmatrix}, \Lambda_C = \begin{bmatrix} 1 & -2 \\ 1 & 0 \\ 1 & 2 \end{bmatrix}.$$

Parameter	Model					
	A		B		C	
	Estimate	SE	Estimate	SE	Estimate	SE
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Age 7	12.094	4.880	12.094	4.880	12.094	4.880
Age 9	57.168	15.992	57.168	15.992	57.168	15.992
Model $T_{ML}(df = 1)$	2.131		2.131		2.131	

Piecewise LCM

$$Y_{it} = \alpha_{iy} + \beta 1_{iy} \lambda_{1t} + \beta 2_{iy} \lambda_{2t} + \varepsilon_{yit}$$

$$\alpha_{iy} = \mu_{\alpha} + \zeta_{\alpha yi}$$

$$\beta 1_{iy} = \mu_{\beta 1} + \zeta_{\beta 1 yi} \quad \Phi_y = \begin{bmatrix} \psi_{\alpha\alpha y} & & \\ \psi_{\alpha\beta 1y} & \psi_{\beta 1\beta 1y} & \\ \psi_{\alpha\beta 2y} & \psi_{\beta 1\beta 2y} & \psi_{\beta 2\beta 2y} \end{bmatrix}$$

$$\beta 2_{iy} = \mu_{\beta 2} + \zeta_{\beta 2 yi}$$

Requires 5 time points,
with at least 2 before and
after the turning point.

$$Y_{it} = \alpha_{iy} + \beta 1_{iy} \lambda_{1t} + \beta 2_{iy} \lambda_{2t} + \varepsilon_{yit}$$

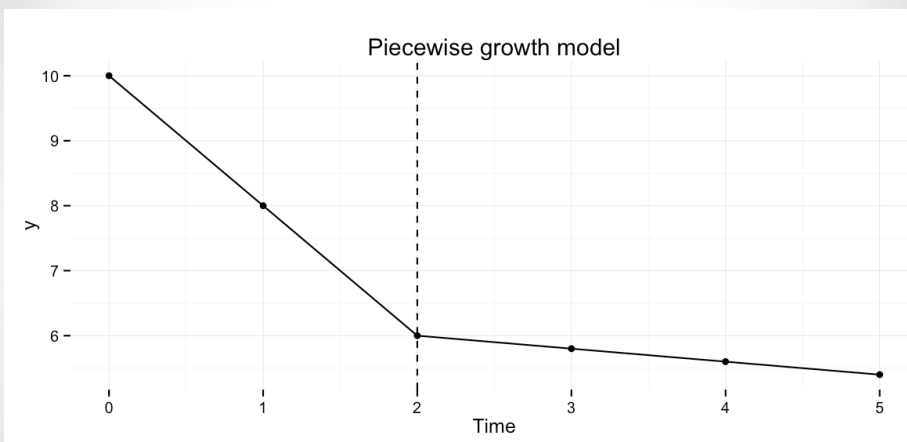
Time codes (λ): Reflects the passage of time.

Linear: 0-1-2-3-4-5 (or some other reflecting the true passage of time, e.g. 0-1-3-4-6).

Quadratic: Addition of a second slope factor with squared time codes (0-1-4-9-16-25).

Piecewise: Two slopes with loadings reflecting the transition point:

- 0-1-2-3-3-3
- 0-0-0-0-1-2-3



Piecewise time codes

- **Approach 1:**

0-1-2-3-3-3-3: The first slope reflects change up to the 4th measurement occasion. Linear growth.

0-0-0-0-1-2-3: The second slope reflect change occurring after the 4th measurement occasion. Linear growth.

- **Approach 2: Added Rate model**

0-1-2-3-4-5-6: The first slope reflects a linear rate of change occurring throughout the study. Linear Growth

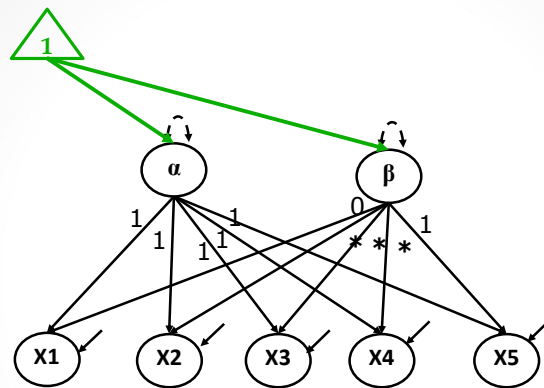
0-0-0-0-1-2-3: The second slope reflect the difference in the rate of change occurring between the two periods (before and after after the 4th measurement occasion). Additive growth.

Turning Points:

- Piecewise modeling is useful when the turning point is known a priori (e.g., intervention, change, retirement, transition, puberty).
- The turning point needs to be common to all participants (not random).
- When the turning point is not known a priori but suspected to exist (e.g., imagine you want to know how long, on average, after retirement it takes for people to change their life rhythm), then it is possible to estimate a fully linear trajectory and to examine the modification indices associated with the loadings on the slope factor to locate the turning point (see Kwok et al., 2010).
- Convergence problems are associated with having only 2 measurement occasions before, but not after, the turning point (Diallo et al., 2015)

Latent Basis / Multibase Models

- So far, we have seen LCM in which all factor loadings representing the time codes are fixed to specific values.
- For identification purposes, only two of them need to be fixed to the pre-specified values of 0 and, typically, 1.
- As always, zero serves to locate the intercept.
- The slope factor will then reflect the total amount of change occurring between the 0 and 1 time points.



$$Y_{it} = \alpha_{iy} + \beta_{iy}\lambda_t + \varepsilon_{yit}$$

$$\alpha_{iy} = \mu_{\alpha} + \zeta_{\alpha yi} \quad \beta_{iy} = \mu_{\beta} + \zeta_{\beta yi}$$

Freely estimated time codes

0 * * * * 1:

Here the slope reflects the total amount of change occurring between the first and last time point (over the course of the study).

The freely estimated unstandardized loadings reflect the proportion of this change that has occurred at each time point.
e.g., 0 .5 .6 .7 .8 1: 50% of the change has occurred by Time 2, 60% by Time 3, etc.

Imagine a slope of 100 and an intercept of 0.

e.g., 0 .5 1.5 .9 .8 1: Non-linear trajectory. 50% of the change has occurred by Time 2 (50), 150% by Time 3 (150), then levels go down by Time 4 (80) and 5 (90), and then up again to 100 by the end.

Freely estimated time codes

0 1 * * * * :

Here the slope reflects the total amount of change occurring between the first and second time point.

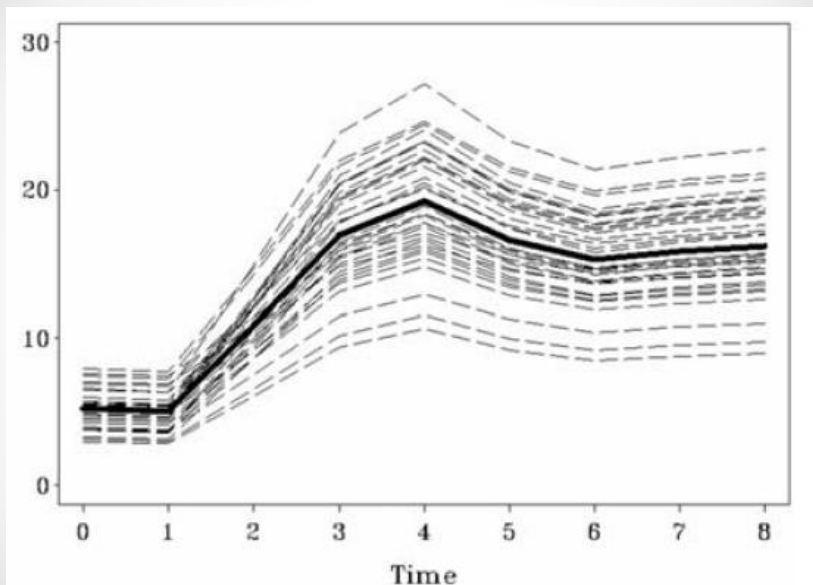
The freely estimated unstandardized loadings reflect how much change (in relation to this initial change level) occurs at each subsequent time point.

Imagine a slope of 10 and an intercept of 0.

e.g., 0 1 2 2.5 2 3 : By time 3, twice the amount of change that has occurred by Time 1 has occurred (20), then 25, then down to 20, then 30.

Limitations of multibase models

- As in all LCM seen so far, the intercept and slope(s) factors are random parameters, free to differ across participants (each participant has his or her own linear, quadratic, or piecewise trajectory).
- In multibase models:
 - The variance of the intercept factor reflects inter-individual variations in initial levels.
 - The variance of the slope factor reflects inter-individual variations in the total amount of change.
 - The loadings are the same for everyone: Each participant's trajectories have the exact same shape.



Example: Linear

Five measurement points, observed (non latent) equally repeated measures (X1 to X5)

Model:

```
i s | X1@0 X2@1 X3@2 X4@3 X5@4 ;  
X1 X2 X3 X4 X5;
```

The | symbol is a shortcut to define a latent curve model.

The **I S** labels can be changed as need. With two of them, a linear LCM will be estimated.

Without the shortcut, the full input would be:

Model:

```
i BY X1@1 X2@1 X3@1 X4@1 X5@1 ;  
s BY X1@0 X2@1 X3@2 X4@3 X5@4 ;  
i s ;  
i with s ;  
[i s];  
[X1@0 X2@0 X3@0 X4@0 X5@0];  
X1 X2 X3 X4 X5;
```

Example: Quadratic

Model:

i s q | X1@0 X2@1 X3@2 X4@3 X5@4 ;
X1 X2 X3 X4 X5;

OR:

Model:

i BY X1@1 X2@1 X3@1 X4@1 X5@1;
s BY X1@0 X2@1 X3@2 X4@3 X5@4 ;
q BY X1@0 X2@1 X3@4 X4@9 X5@16 ;
i s q; i with s q; s with q;
[i s q];
[X1@0 X2@0 X3@0 X4@0 X5@0];
X1 X2 X3 X4 X5;

Example: Piecewise

Model:

i s1 | X1@0 X2@1 X3@2 X4@2 X5@2 ;
i s2 | X1@0 X2@0 X3@0 X4@1 X5@2 ;
X1 X2 X3 X4 X5;

OR:

Model:

i BY X1@1 X2@1 X3@1 X4@1 X5@1;
s1 BY X1@0 X2@1 X3@2 X4@2 X5@2 ;
s2 BY X1@0 X2@0 X3@0 X4@1 X5@2 ;
i s1 s2; i with s1 s2; s1 with s2;
[i s1 s2]; [X1@0 X2@0 X3@0 X4@0 X5@0];
X1 X2 X3 X4 X5;

Example: Multibase

MODEL:

i s | X1@0 X2* X3* X4* X5@1 ;

X1 X2 X3 X4 X5;

Or

i s | X1@0 X2@1 X3* X4* X5 * ;

X1 X2 X3 X4 X5;

Or

Model:

i BY X1@1 X2@1 X3@1 X4@1 X5@1;

s BY X1@0 X2@1 X3* X4* X5* ;

i s ; i with s;

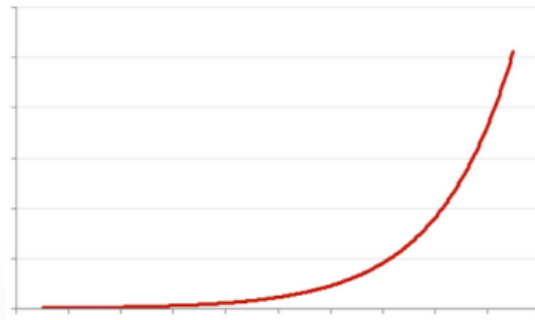
[i s];[X1@0 X2@0 X3@0 X4@0 X5@0];

X1 X2 X3 X4 X5;

Other non-linear trajectories

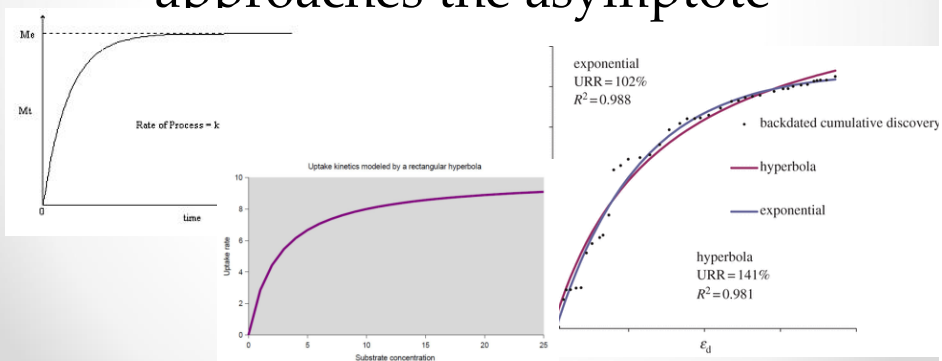
- Complex.
- Not often useful.
- Often include non-random parameters (that do not vary across individuals, and yet should).
- Exponential, Logistic, negative exponential, rectangular hyperbolic, Gompertz.
- See:

Exponential: Accelerating rate of change



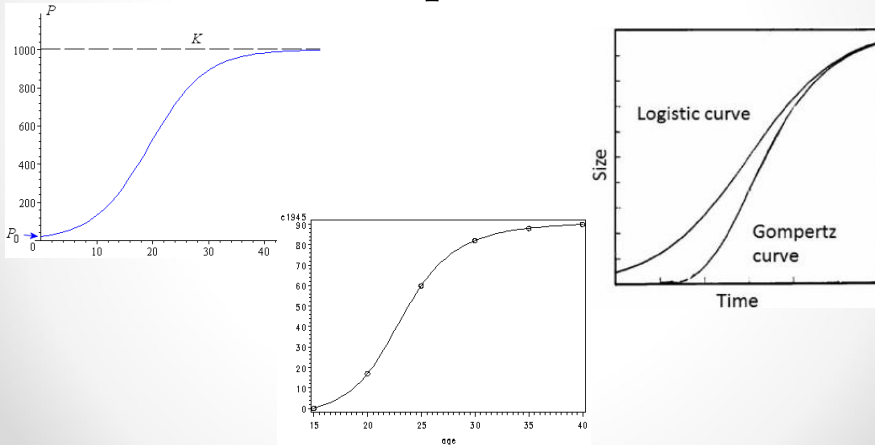
Negative Exponential / Rectangular Hyperbolic:

- * Quick growth reaching a ceiling
- * Rate of change that increases as one approaches the asymptote



Logistic/ Gompertz:

* Difference: The logistic is symmetric, the Gompertz is not



See:

- Wu, W., Selig, J.P., & Little, T.D. (2013). Longitudinal Data analysis. In T.D. Little (Ed.) *The Oxford Handbook of Quantitative Methods in Psychology, Volume 2* (pp. 387-410). New York: Oxford University Press.
- Hancock, G.R., Harring, J.R., & Lawrence, F.R. (2013). Using latent growth modelign to evaluate longitudinal change. In G. R. Hancock & R. O. Mueller (Eds), *Structural Equation Modeling: A Second Course, 2nd edition* (pp. 309-342). Greenwich, CO: IAP.
- Ram, N., & Grimm, K. (2007). Using simple and complex growth models to articulate developmental change: Matching theory to method. *International Journal of Behavioral Development*, 31, 303-316.
- Grimm, K.J., & Ram, N. (2009). Nonlinear growth models in Mplus and SAS. *Structural Equation Modeling*, 16, 676-701.
- Preacher, K.J., & Hancock, G.R. (2015). Meaningful aspects of change as novel random coefficients: A general method for reparametrizing longitudinal models. *Psychological Methods*, 20, 84-101.

Conditional Models

- **Time Invariant Covariates**

Incorporation of predictors that are either only measured at the beginning of the study, or outcomes that are only measured at the end of the study, or of predictors/outcomes that do not change over time (sex, etc.).

Typically, these are only allowed to predict, or be predicted, by the intercept and slope(s) factors.

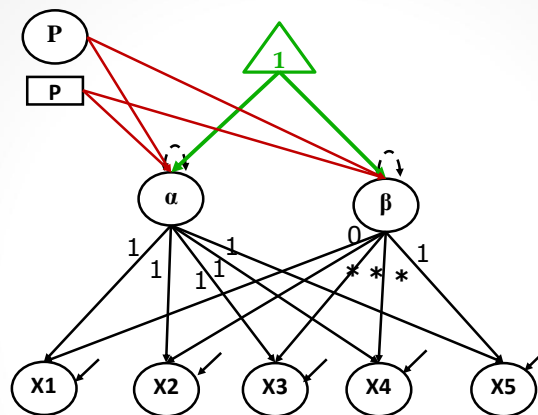
- **Time Varying Covariates**

Incorporation of predictors/outcomes that differ across time points.

Typically, these are only allowed to predict, or be predicted, but the time-varying observations. The invariance over time of these predictions can be tested.

- **Multiple Processes**

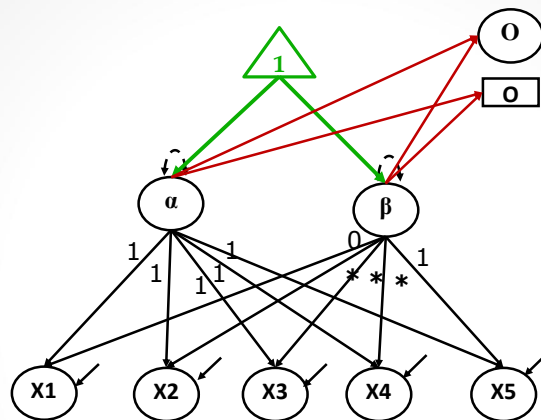
Incorporation of more than one LCM in the same model.



$$Y_{it} = \alpha_{iy} + \beta_{iy}\lambda_t + \varepsilon_{yit}$$

$$\alpha_{iy} = \mu_{\alpha} + \gamma_{\alpha 1}x_{1i} + \zeta_{\alpha yi}$$

$$\beta_{iy} = \mu_{\beta} + \gamma_{\beta 1}x_{1i} + \zeta_{\beta yi}$$



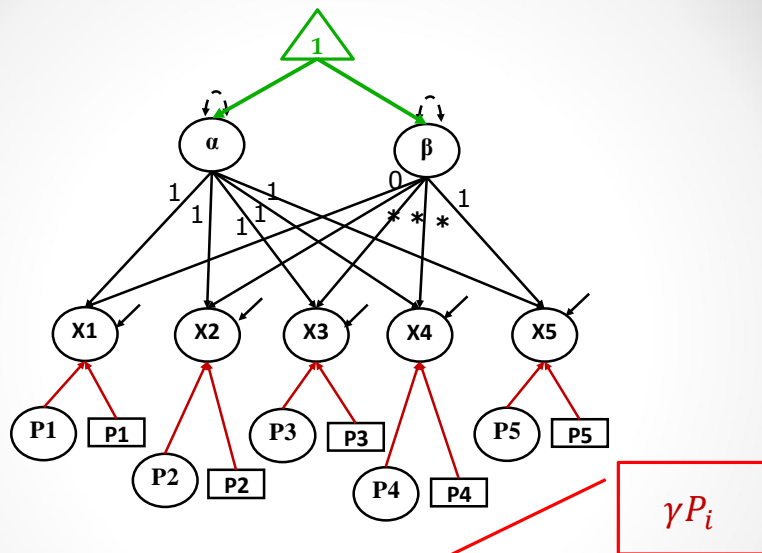
$$O_i = \alpha_0 + \beta_0 \alpha_{iy} + \beta_0 \beta_{iy} \epsilon_{oi}$$

Model:

i BY X1@1 X2@1 X3@1 X4@1 X5@1;
s BY X1@0 X2@1 X3@2 X4@3 X5@4 ;
i s ; i with s;
[i s]; [X1@0 X2@0 X3@0 X4@0 X5@0];
X1 X2 X3 X4 X5;
i s ON PI P2;

Model:

i BY X1@1 X2@1 X3@1 X4@1 X5@1;
s BY X1@0 X2@1 X3@2 X4@3 X5@4 ;
i s ; i with s;
[i s]; [X1@0 X2@0 X3@0 X4@0 X5@0];
X1 X2 X3 X4 X5;
O1 O2 ON i s;



$$Y_{it} = \alpha_{iy} + \beta_{iy}\lambda_t + \gamma_t P_{it} + \varepsilon_{yit}$$

$$\alpha_{iy} = \mu_\alpha + \zeta_{\alpha yi} \qquad \beta_{iy} = \mu_\beta + \zeta_{\beta yi}$$

Model:

i BY X1@1 X2@1 X3@1 X4@1 X5@1;

s BY X1@0 X2@1 X3@2 X4@3 X5@4 ;

i s ; i with s;

[i s]; [X1@0 X2@0 X3@0 X4@0 X5@0];

X1 X2 X3 X4 X5;

X1 ON P1;

X2 ON P2;

X3 ON P3;

X4 ON P4;

X5 ON P5;

X1-X5 PON P1-P5;

Model:

i BY X1@1 X2@1 X3@1 X4@1 X5@1;

s BY X1@0 X2@1 X3@2 X4@3 X5@4 ;

i s ; i with s;

[i s]; [X1@0 X2@0 X3@0 X4@0 X5@0];

X1 X2 X3 X4 X5;

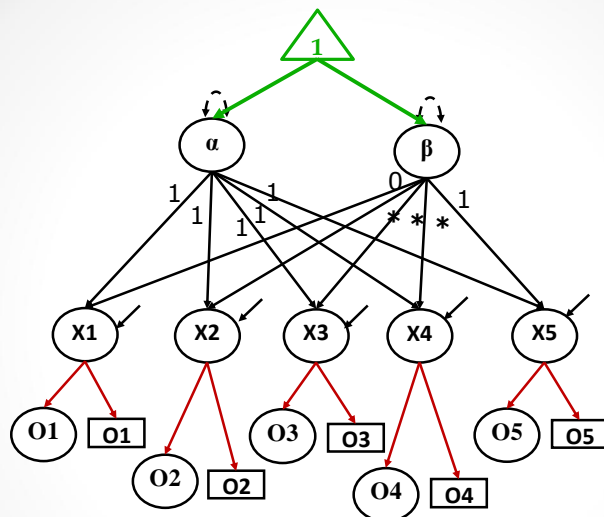
X1 ON P1 (a);

X2 ON P2 (a);

X3 ON P3 (a);

X4 ON P4 (a);

X5 ON P5 (a);



$$O_{it} = \alpha_{O_t} + \rho_{O_t} X_{it} + \varepsilon_{oit}$$

Model:

i BY X1@1 X2@1 X3@1 X4@1 X5@1;

s BY X1@0 X2@1 X3@2 X4@3 X5@4 ;

i s ; i with s;

[i s]; [X1@0 X2@0 X3@0 X4@0 X5@0];

X1 X2 X3 X4 X5;

O1 ON X1;

O2 ON X2;

O1-O5 PON X1-X5;

O3 ON X3;

O4 ON X4;

O5 ON X5;

Model:

i BY X1@1 X2@1 X3@1 X4@1 X5@1;

s BY X1@0 X2@1 X3@2 X4@3 X5@4 ;

i s ; i with s;

[i s]; [X1@0 X2@0 X3@0 X4@0 X5@0];

X1 X2 X3 X4 X5;

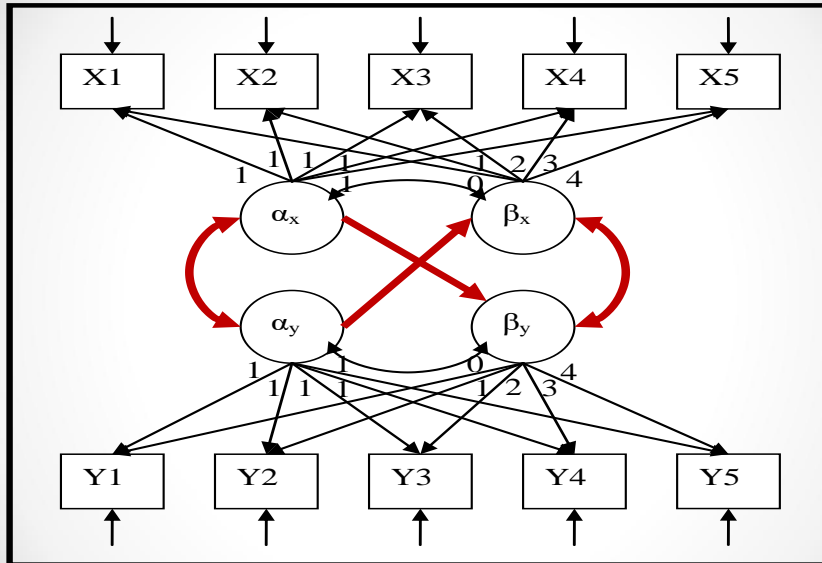
O1 ON X1 (a);

O2 ON X2 (a)

O3 ON X3 (a);

O4 ON X4 (a);

O5 ON X5 (a);



Model:

ix BY X1@1 X2@1 X3@1 X4@1 X5@1;

sx BY X1@0 X2@1 X3@2 X4@3 X5@4 ;

iy BY Y1@1 Y2@1 Y3@1 Y4@1 Y5@1;

sy BY Y1@0 Y2@1 Y3@2 Y4@3 Y5@4 ;

ix sx iy sy ;

[ix sx iy sy];

ix with sx; iy with sy;

ix with iy; sx with sy;

[X1@0 X2@0 X3@0 X4@0 X5@0];

[Y1@0 Y2@0 Y3@0 Y4@0 Y5@0];

X1 X2 X3 X4 X5 Y1 Y2 Y3 Y4 Y5;

sx ON iy;

sy ON ix;

Thank you!